



# Clínica y Salud

<https://journals.copmadrid.org/clysa>



## Psychometric Properties of the Spanish Version of the Cognitive Flexibility Inventory

Irene Jaén<sup>1</sup>, Verónica Vidal-Arenas<sup>1</sup>, Carlos Suso-Ribera<sup>1,2</sup>, M. Carmen Pastor<sup>1</sup>, and Azucena García-Palacios<sup>1,2</sup>

<sup>1</sup>Universitat Jaume I, Castelló de la Plana, Spain; <sup>2</sup>Instituto de Salud Carlos III, Madrid, Spain

### ARTICLE INFO

#### Article history:

Received 4 July 2023  
Accepted 20 October 2023  
Available online 23 April 2024

#### Keywords:

Cognitive Flexibility  
Internal consistency  
Construct validity  
Adaptation

### ABSTRACT

**Objective:** Cognitive flexibility, identified as a transdiagnostic process in psychopathology, lacks specific Spanish-validated assessment tools. This study aims to examine the psychometric properties of the Cognitive Flexibility Inventory in a Spanish sample ( $N = 300$ , mean age = 22.66,  $SD = 4.92$ ), 76% women). **Method:** An exploratory structural equation model (ESEM) was used to test the inventory's latent structure. Reliability was calculated, as well as construct validity evidence with measures of emotion regulation, attributional style, and depression. **Results:** The original bifactorial model of the Cognitive Flexibility Inventory (Alternatives and Control scales) was replicated and obtained good fit indices. One item was excluded due to problematic factor loadings, yielding a final version with 19 items. Evidence of validity and reliability was obtained. **Conclusions:** The Spanish version of the Cognitive Flexibility Inventory showed satisfactory psychometric properties, making it a suitable measure for the assessment of cognitive flexibility in the Spanish population.

### Las propiedades psicométricas de la versión española del Inventario de Flexibilidad Cognitiva

### RESUMEN

**Objetivo:** La flexibilidad cognitiva, identificada como un proceso transdiagnóstico en la psicopatología, carece de herramientas específicas validadas en español. Este estudio se propone examinar las propiedades psicométricas del Inventario de Flexibilidad Cognitiva en una muestra española ( $N = 300$ , media de edad = 22.66,  $DT = 4.92$ ), 76% mujeres). **Método:** Se utilizó un modelo exploratorio de ecuaciones estructurales para comprobar la estructura latente. Se calculó la fiabilidad, así como la evidencia de validez de constructo con medidas de regulación emocional, estilo atribucional y depresión. **Resultados:** Se replicó el modelo bifactorial original del Inventario de Flexibilidad Cognitiva (escalas de alternativas y control) y se obtuvieron buenos índices de ajuste. Se excluyó un ítem debido a cargas factoriales problemáticas, por lo que la versión final incluyó 19 ítems. Se obtuvieron pruebas de validez y fiabilidad. **Conclusiones:** La versión española del Inventario de Flexibilidad Cognitiva mostró propiedades psicométricas satisfactorias, por lo que es una medida adecuada para la evaluación de la flexibilidad cognitiva en la población española.

Cognitive flexibility (CF) has been defined by many authors as the ability that enables people to effectively adjust their cognitive process to changing environmental demands (Deák, 2003; Dennis & Vander Wal, 2010). Thus, high cognitive flexibility is essential to successfully challenge and restructure maladaptive beliefs into a more adaptive thinking style (Dennis & Vander Wal, 2010). In recent years, this ability has been considered an essential mechanism associated with mental health (Bonanno & Burton, 2013; Kashdan & Rottenberg, 2010; Wersebe et al., 2018). However, one of the major problems of this construct has been the lack of consensus about its definition and measurement. For example, the same term has been used to refer to

different constructs due to the different psychological traditions that consider cognitive flexibility a focus of study (Ionescu, 2012, 2017).

A first example of that is the definition of CF as a cognitive mechanism of executive functioning (Geurts et al., 2009; Miyake et al., 2000). According to this view, CF has been generally measured using neurocognitive tests, such as the Wisconsin Card Sorting Test, the Stroop Colour Test, or the Trail Making Test, which assess the flexible/perseverative response pattern of an individual when conducting tasks that require shifting mental sets in response to stimuli. In addition to these, experimental psychologists have also developed task-switching variants of these neuropsychological

Cite this article as: Jaén, I., Vidal-Arenas, V., Suso-Ribera, C., Pastor, M. C., & García-Palacios, A. (2024). Psychometric properties of the Spanish version of the Cognitive Flexibility Inventory. *Clínica y Salud*. Ahead of print. <https://doi.org/10.5093/clysa2024a3>

Funding: This study was supported by the Universitat Jaume I, under the Grant UJI-B2019-33 and a Predoctoral Grant PREDOC/2017/26. Correspondence: [ijaen@uji.es](mailto:ijaen@uji.es) (I. Jaén)

ISSN: 1130-5274/© 2024 Colegio Oficial de la Psicología de Madrid. This is an open access article under the CC BY-NC-ND license (<http://creativecommons.org/licenses/by-nc-nd/4.0/>).

tests that involve switching between different mental operations in response to task cues (Allport et al., 1994; Rogers & Monsell, 1995; Waugh et al., 2011).

Another conceptualization of CF refers to the rigidity or the ability to adapt our thoughts and behaviours to better respond to the context (Deák, 2003; Garcia-Garcia et al., 2010). In fact, CF is a core mechanism of change in therapies such as the Cognitive Behavioural Therapy (Beck et al., 1979), as well as in more modern, transdiagnostic treatments, such as the Unified Protocol (Boisseau et al., 2010), which propose that CF would be a transdiagnostic process underlying psychopathology. When assessing this conceptualization of CF, authors usually rely on self-report measures, such as the Cognitive Flexibility Scale (CFS; M. M. Martin & Rubin, 1995) and the Cognitive Flexibility Inventory (Dennis & Vander Wal, 2010), which have become the most popular. Compared to the majority of neuropsychological tests, these self-report measures are significantly faster to administer and score and are less sensitive to practice effects (McCaffrey et al., 1992). This view of CF as an ability to flexibly change one's thoughts and obtain different interpretations of a situation when needed is also more consistent with the perspective of modern psychological treatments (Boisseau et al., 2010).

The CFS is one of the first self-report measures of CF. The CFS evaluates CF linked to interpersonal communication competence, so the authors proposed that cognitive flexibility should be considered an essential component of this interpersonal ability (M. M. Martin & Anderson, 1998; M. M. Martin & Rubin, 1995). This, however, has been argued to be a limitation to the study of CF since it is not clear whether communication flexibility can be compared to the CF required to change and replace maladaptive cognitions effectively (Dennis & Vander Wal, 2010). As a consequence, Dennis and Vander Wal (2010) developed the Cognitive Flexibility Inventory, an alternative measure of CF that can be administered for research and clinical purposes to evaluate an individual's ability to think adaptively and flexibly when facing stressful life events.

The Cognitive Flexibility Inventory (Dennis & Vander Wal, 2010) is a brief self-report measure composed of two subscales: Alternatives and Control. The Alternatives subscale measures the tendency to perceive multiple alternative explanations of life events and generate multiple solutions to difficult situations (e.g., "I consider multiple options before making a decision"), whereas the Control subscale was developed to test the tendency to perceive difficult situations as controllable (e.g., "I feel I have no power to change things in difficult situations"). The original English scale (Dennis & Vander Wal, 2010) supported this two-factor structure and had good evidence of internal consistency ( $> .84$ ) and test-retest reliability ( $r = .81$ ). The scale has now been adapted into Chinese (Wang et al., 2016), Japanese (Oshiro et al., 2016), Iranian (Shareh et al., 2014), Russian (Kurginyan & Osavolyuk, 2018), and Italian (Portoghese et al., 2020). However, some inconsistencies have been found in the latent structure in the cross-cultural adaptations. For example, a three-factor structure (i.e., Control, Alternatives, and Alternatives for human behaviours) was reported in the Iranian version (Shareh et al., 2014). In addition, while the Russian (Kurginyan & Osavolyuk, 2018) and the Italian (Portoghese et al., 2020) adaptations replicated the original two-factor structure of the scale, they differed in the number of items that composed the two-factor structure. Specifically, item 1 ("I am good at 'sizing up' situations") was problematic in both adaptations. This item loaded into the Control subscale in the Russian version (Kurginyan & Osavolyuk, 2018), while it had high cross-loading into both factors in the Italian version, which led the authors to remove this item from their adaptation of the inventory (Portoghese et al., 2020).

Regarding the construct validity of the Cognitive Flexibility Inventory according to external criteria, the different versions of the inventory have shown that CF is positively associated with the use of more adaptive coping efforts and less maladaptive

strategies (Kurginyan & Osavolyuk, 2018). In addition, research with the Cognitive Flexibility Inventory has indicated that depressed individuals perform poorly in CF (Dennis & Vander Wal, 2010; Shareh et al., 2014), which supports the idea that CF might be an underlying mechanism in emotional disorders such as depression (Boisseau et al., 2010). Also, Demirtas (2020) showed that cognitive flexibility positively correlated with self-efficacy and mental well-being.

Research has also tested the construct validity of the Cognitive Flexibility Inventory in relation to other measures of CF, such as the Attributional Styles Questionnaire (ASQ; Peterson et al., 1982). Surprisingly, though, these studies showed a negative association between the Cognitive Flexibility Inventory and the ASQ, thus suggesting that greater CF on the Cognitive Flexibility Inventory would be related to less cognitive flexibility on the ASQ (Dennis & Vander Wal, 2010; Kurginyan & Osavolyuk, 2018). In this regard, some studies have used the ASQ as a measure of CF using the intra-individual standard deviation for the stable and global items associated with the six negative events (Fresco et al. 2017). However, previous validations considered extreme scores as indicative of less CF for the operationalization of CF (Dennis & Vander Wal, 2010; Kurginyan & Osavolyuk, 2018), which might not be adequate as a measure of CF and might explain the surprising findings reported in past research.

Although evidence of construct validity and reliability of the Cognitive Flexibility Inventory has been studied in several languages, an adaptation for the Spanish population is still missing. The present study will adapt the Cognitive Flexibility Inventory into Spanish and examine its psychometric properties in a Spanish sample to provide a self-report measure to evaluate the flexibility required to successfully modify and restructure maladaptive beliefs with a more adaptive thinking. Because there is no evidence about the structure of the Cognitive Flexibility Inventory in the Spanish population and considering the aforementioned inconsistencies in the internal structure of this measure in prior cross-cultural validations, an exploratory structural equation modelling (ESEM) will be used to investigate the internal structure of the scale. In addition, sources of construct validity of the scale, together with estimates of its internal consistency and stability over time (test-retest), will be examined. Empirical evidence supporting the two-factor solution obtained in the original version of the scale (Dennis & Vander Wal, 2010) is expected to be found. Also, significant positive associations (i.e., convergent construct validity) between the Cognitive Flexibility Inventory and the use of adaptive coping strategies and favorable attributional styles (a construct closely related to flexibility in cognitive thinking) are expected. Lastly, a negative relationship between CF and depressive symptomatology, in line with previous research is expected.

## Method

### Participants

Participants were 300 undergraduate students from Spain (mean age = 22.66,  $SD = 4.92$ , 76% females). The sample size was determined according to the graded scale of sample sizes for scale development proposed by Guadagnoli & Velicer (1988), which recommends a minimum of 300 participants for obtaining satisfactory results in scale validation. The study was advertised using printed posters at the author's university. Once the participants contacted the researchers by email, they were provided with the web link to a Qualtrics survey with all the questionnaires. Participants completed all the measures on the same administration.

The Cognitive Flexibility Inventory was administered again one month after the initial administration to assess test-retest reliability. This second administration was completed by a subsample of 85 participants (mean age = 23.81,  $SD = 4.81$ , 80% females,  $n = 68$ ), who

had visited the laboratory to participate in another study where they completed the CFI first, followed by other measures. As a token of appreciation for their participation in the study, all participants received a compensation of 10 euros. In addition, participants who completed the second assessment received an additional 10 euros.

The present study was approved by the ethical committee of the Jaume I University (CD/37/2021) and it was performed in accordance with the Declaration of Helsinki (1964). All the participants gave their written informed consent to participate in the study.

## Instruments

### **Spanish Version of the Cognitive Flexibility Inventory** (Dennis & Vander Wal, 2010; Appendix A and B)

This is a 20-item questionnaire composed of two subscales: Alternatives (13 items) and Control (7 items). The Alternatives subscale evaluates the ability to find multiple explanations for a problem and to produce different solutions in front of difficult situations. The Control subscale measures the tendency to perceive situations as controllable. Items are scored using a 7-point Likert scale, which ranges from 1 (*strongly disagree*) to 7 (*strongly agree*). Higher scores in both scales reflect greater CF. Internal consistency for the Alternatives and Control subscales in previous research were .86 and .91, respectively (Dennis & Vander Wal, 2010). The translation into the Spanish language was conducted following the back translation method (Brislin, 1973). First, two of the authors of the present study, who are psychologists and native Spanish speakers and proficient in English, translated the English version into Spanish. These authors were familiar with the target language and culture and intent of the scale. Second, a native-English-speaking professional translator who was not familiar with the questionnaire back-translated the questionnaire from Spanish to English. Then, the original and back-translated versions were compared, with a focus on appropriateness, clarity, and meaning of each item. Any discrepancies identified between the original and the back-translated items were discussed and resolved between all the study authors. For example, in item 1, the term “sizing up” was translated as “*evaluando*” which equates to “evaluating” in English. Likewise, in item 9, the word “troublesome” was translated as “*problemático*”, which means “problematic” in English.

### **The Attributional Styles Questionnaire** (ASQ; Peterson et al., 1982; Sanjuán & Magallares, 2008)

This questionnaire consists of twelve hypothetical life events (6 positive and 6 negative). Individuals are asked to describe a plausible cause if these events happened to them. These causes are rated on a 7-point scale following these anchors: (1) internal or external, (2) stable or unstable, and (3) global or specific. Two composite indexes can be obtained based on the scores of the positive and negative situations, one for positive attributional style and one for negative attributional style, respectively. Reliability coefficients (Cronbach's alpha) in the present study for positive and negative events were .81 and .74, respectively.

### **The Cognitive Emotion Regulation Questionnaire** (CERQ; Dominguez-Sánchez et al., 2013; Garnefski et al., 2001)

It is a 36-item scale that evaluates the cognitive aspects of emotion regulation. Specifically, the questionnaire comprises 9 subscales, each consisting of different ways in which an individual can think after experiencing threatening or stressful events: Acceptance, Positive reappraisal, Positive refocusing, Putting into perspective, Catastrophizing, Self-blame, Other-blame,

Rumination, and Refocusing on planning. Each scale consists of four items measured on a 5-point scale, ranging from 1 (*almost never*) to 5 (*almost always*). Higher scores indicate more frequent use of that strategy. Reliability coefficients (Cronbach's alpha) of the nine CERQ subscales in this study ranged from .62 to .88.

### **The Beck Depression Inventory-II** (BDI-II; Beck et al., 1996; Sanz et al., 2003)

This inventory measures the severity of the depression symptomatology during the last week. The BDI-II is composed of 21 items that reflect cognitive, behavioural, emotional, and somatic symptoms common to depression. Items are rated on a 3-point scale. Higher scores indicate higher severity of the depression symptomatology. The Cronbach's alpha of the scale in the present study was .92.

## Data Analysis

To investigate the internal structure of the Cognitive Flexibility Inventory we used an exploratory structural equation modeling (ESEM) analysis with two latent factors (to replicate the original structure) using Mplus 8.4. We used a weighted least square mean and variance adjusted (WLSMV) estimator due to the non-normal distribution of the data. To investigate the fit of the data to the models, we used the comparative fit indexes (CFI), the Tucker-Lewis index (TLI), and the root mean square error of approximation (RMSEA) indices. For interpretation purposes, CFI and TLI >.90 and >.95 indicate acceptable and optimal fits, respectively (Marsh et al., 2004). RMSEA values  $\leq$ .10 indicate an acceptable fit (Weston & Gore, 2006). Confidence intervals (95%) were also calculated and reported. Cronbach's alpha and omega were used to examine the internal consistency of our adaptation of the Cognitive Flexibility Inventory. Cohen's *d* was calculated to compare the effect sizes of the means obtained in each subscale by sex. In addition, differential item functioning (DIF) was calculated with the purpose of detecting whether items of the Cognitive Flexibility Inventory function differently depending on sex. To test for differential item functioning (DIF) in responses based on sex, we followed a series of steps to calculate a single covariate MIMIC model. These steps included: (1) confirmatory factor analysis (CFA) for the total sample, (2) building the MIMIC model without direct effects, and (3) if the modification indices suggested significant direct effects, the model was further tested with these additions, following the approach proposed by Jones (2006). Also, intra-class correlation coefficients (ICC; CI 95%) were used to explore the temporal stability of the Cognitive Flexibility Inventory in a one-month retest. Finally, sources of construct validity were analysed by computing Pearson correlations between the Cognitive Flexibility Inventory and measures of cognitive emotion regulation (CERQ), attributional styles (ASQ), and depression symptomatology (BDI-II). According to Cohen (1992), correlation values  $\geq$ .10, .30 and .50 are considered small, medium, and large effect sizes, respectively. Reliability and construct validity analyses were conducted using SPSS v.26. The datasets generated during and/or analysed during the current study are available in [https://osf.io/3fbx5/?view\\_only=119640dc32bf47b2b508718f180cf165](https://osf.io/3fbx5/?view_only=119640dc32bf47b2b508718f180cf165).

## Results

### **Descriptive Analysis of the Items**

The distributional properties of the 20 items of the Spanish version of the Cognitive Flexibility Inventory (i.e., means, standard deviations, skewness, and kurtosis) are presented in Table 1. Specifici-

cally, skewness ranged from -1.41 to 0.58, and kurtosis ranged from -1.20 to 3.53. Thus, we obtained asymmetric values for items 3, 13, 15 and 18. Item 3 revealed a leptokurtic distribution.

**Table 1.** Descriptive Analysis of the Items of the Spanish Version of the Cognitive Flexibility Inventory

	Mean	SD	Skewness	Kurtosis
Item 1	5.47	1.08	-0.86	1.16
Item 2	4.14	1.77	0.01	-1.20
Item 3	5.91	1.04	-1.41	3.53
Item 4	3.69	1.72	0.31	-1.13
Item 5	5.52	1.14	-0.62	-0.13
Item 6	4.67	1.41	-0.42	-0.20
Item 7	3.37	1.74	0.43	-1.00
Item 8	6.16	1.42	-0.99	0.83
Item 9	3.24	1.56	0.35	-0.70
Item 10	5.80	1.15	-1.00	0.86
Item 11	3.29	1.59	0.56	-0.78
Item 12	6.30	0.78	-0.84	-0.08
Item 13	5.56	1.18	-1.11	1.54
Item 14	5.60	1.19	-0.91	0.63
Item 15	5.58	1.14	-1.16	1.62
Item 16	5.09	1.12	-0.10	-0.59
Item 17	3.51	1.46	0.34	-0.78
Item 18	5.49	1.15	-1.25	2.00
Item 19	5.48	1.14	-0.95	1.27
Item 20	5.59	1.15	-0.98	1.07

### Structural Validity Evidence

Structure analyses displayed acceptable/optimal fit indices for the tested model (CFI = .938, TLI = .922, RMSEA = .098, CI 90% [.090, .106]). As shown in Table 2, all the items significantly loaded into one of the two factors proposed by the original version of the Cognitive Flexibility Inventory with loadings over .40. Item 15, which originally belonged to the Control scale, had cross-loadings in both

factors, but had a higher loading into the Alternatives factor. Then, a new ESEM analysis was carried out to test the new structure based on 19 items (i.e., without Item 15). Results showed good fit indices (CFI = .948, TLI = .934, RMSEA = .094, CI 90% [.085, .103]), and also salient loadings of each item for corresponding sub-scale (Alternatives, > .39, Control, >.66; see Table 2). Therefore, the following analyses were conducted based on the 19-item structure. Results showed a significant correlation between both factors ( $r = .134$ ).

### Reliability Evidence and Descriptive Coefficients

Cronbach's alpha and omega indices for each factor are presented in Table 3. All coefficients, both in the total sample and across sex, were above .86. Cohen's  $d$  results show no sex differences between the Cognitive Flexibility Inventory subscales. However, significant gender-related effects were observed in the analysis of differential item functioning (DIF). Specifically, DIF by gender was found for item 10 (i.e., "*Soy bueno poniéndome en la piel de otra gente*"), with fit indices CFI = .937, TLI = .922, and RMSEA (90% CI) = .095 [.087, .103], indicating that females scored higher than males on this item. After accounting for this direct effect in the MIMIC model for CFI (CFI = .941, TLI = .927, RMSEA 90% CI = .092 [.084, .100]), no other significant effects were observed. One-month retest correlations ranged from .67 to .90.

### Construct Validity Evidence

To evaluate the construct validity of the Spanish version of the Cognitive Flexibility Inventory, bivariate correlations were conducted with measures of emotion regulation, attributional styles, and depression (Table 4). The results showed that the two subscales were generally positively associated with the use of adaptive cognitive emotion regulation strategies and inversely related to maladaptive forms of emotion regulation. Specifically, the Alternatives subscale was positively associated with Positive refocusing, Positive reappraisal, Putting into perspective, Rumination, and Refocusing, while negatively linked to Catastrophizing, and Other-blame. The

**Table 2.** Standardized Factor Loadings from the Exploratory Structural Equation Model of the Cognitive Flexibility Inventory Assuming a Two-factor Structure

	20-items form				19-items form			
	Alternatives		Control		Alternatives		Control	
	$\Lambda$	95% CI	$\Lambda$	95% CI	$\Lambda$	95% CI	$\Lambda$	95% CI
Item 1	.49	.43, .57	.16	.08, .24	.48	.41, .55	.15	.07, .23
Item 3	.66	.60, .73	-.14	-.22, -.06	.66	.60, .73	-.15	-.23, -.06
Item 5	.69	.63, .74	-.05	-.12, .02	.69	.63, .74	-.05	-.12, .02
Item 6	.54	.47, .60	.13	.05, .21	.55	.48, .61	.13	.05, .21
Item 8	.44	.36, .53	-.09	-.19, -.003	.45	.36, .53	-.10	-.19, .004
Item 10	.40	.30, .50	-.08	-.18, .03	.40	.29, .50	-.08	-.18, .02
Item 12	.59	.51, .67	.02	-.08, .13	.59	.51, .66	.02	-.09, .12
Item 13	.81	.77, .86	-.08	-.14, -.01	.82	.78, .86	-.08	-.14, -.01
Item 14	.86	.82, .90	-.09	-.15, -.03	.86	.83, .90	-.09	-.15, -.03
Item 15	.48	.40, .55	-.10	.01, .19	-	-	-	-
Item 16	.63	.56, .69	.07	-.01, .15	.63	.56, .69	.07	-.01, -.15
Item 18	.79	.75, .83	.06	.000, .13	.78	.74, .83	.06	-.01, .12
Item 19	.76	.71, .82	.11	.03, .18	.76	.71, .81	.10	.03, .18
Item 20	.86	.83, .90	.04	-.01, .09	.86	.83, .90	.04	.02, .09
Item 2	-.04	-.11, .02	.76	.71, .81	-.03	-.10, .03	.76	.71, .81
Item 4	.02	-.04, .07	.82	.78, .86	.03	-.03, .08	.82	.77, .86
Item 7	.05	-.01, .12	.77	.72, .82	.06	-.01, .13	.77	.72, .82
Item 9	-.10	-.18, -.02	.73	.67, .78	-.09	-.17, -.01	.73	.67, .79
Item 11	.04	.02, .09	.90	.87, .93	.05	-.001, .11	.90	.87, .93
Item 17	.02	-.09, .05	.67	.61, .73	-.02	-.09, .05	.67	.61, .73

Control subscale was positively associated with Positive reappraisal, Positive refocusing, and refocusing on planning, while negatively associated with Rumination, Catastrophizing and Self-blame.

Finally, both the Alternatives and the Control subscales were positively associated with the positive-composite attributional style but negatively linked to the negative-composite attributional style and the severity of depression symptoms. The strength of the correlations was generally small, except for a moderate correlation between the Alternatives scale and the CERQ-RP.

**Table 3.** Descriptive for Females and Males and Reliability Coefficients

	Cronbach's Alpha (95% CI)	Omega (95% CI)	Test-Retest ICC (95% CI)	Mean score (SD)	Cohen's d (a-b)
<b>Alternatives</b>					
Total	.87 (.86,.90)	.88 (.86, .90)	.85 (.77, .90)	77.21 (10.07)	
<sup>a</sup> Females	.89 (.86,.91)	.89 (.86, .91)	.86 (.77, .91)	77.09 (10.30)	.05
<sup>b</sup> Males	.86 (.80,.90)	.86 (.78, .90)	.67 (.07, .88)	77.61 (9.36)	
<b>Control</b>					
Total	.89 (.87,.91)	.89 (.88, .91)	.90 (.85, .94)	22.55 (8.60)	
<sup>a</sup> Females	.89 (.86,.91)	.89 (.87, .91)	.90 (.84, .94)	22.68 (8.59)	.06
<sup>b</sup> Males	.91 (.87,.94)	.91 (.87, .94)	.90 (.71, .96)	22.18 (8.70)	

## Discussion

This study aimed to adapt the Cognitive Flexibility Inventory into the Spanish language and examine its psychometric properties in a sample of Spaniards. Following the original developments by Dennis and Vander Wal (2010), we found an acceptable model fit for a two-factor structure of the inventory. However, our analyses suggested a modified two-factor version composed of 14 items in the Alternatives subscale and 6 items in the Control subscale. Similar to previous adaptations of the Cognitive Flexibility Inventory (Kurginyan & Osavolyuk, 2018), a problematic item in the scale was found. Specifically, item 15 (“I am capable of overcoming the difficulties in the life that I face”) loaded on the Alternatives factor as opposed to the Control factor, which is inconsistent with the distribution proposed in the original English version (Dennis & Vander Wal, 2010). After a discussion about the meaning of Item 15 in the Spanish version and the theoretical implications of changing the structure of the Cognitive Flexibility Inventory based on analytical evidence only, the authors of the present study removed this item

from the scale. Consequently, we obtained a 19-item version of the Cognitive Flexibility Inventory for Spanish speakers, which showed better fit indices than the 20-item form, as well as reliability and validity evidence.

It is important to note that item 15 is the only direct item that originally belonged to the Control subscale, whereas all items in the Alternative subscale are direct. Evidence has suggested that there is a possible acquiescence bias when items are positively worded, which can affect the reliability and dimensionality of the test by secondary sources of variance when combining direct and reverse items (Checa & Espejo, 2018; Suárez-Alvarez et al., 2018). In fact, the inclusion of positive and negative items when measuring two poles of similar constructs may result in a two-factor structure in which positive and negative items load on different factors (Brown, 2003; Spector et al., 1997). In this regard, it would be advisable that future studies explore the psychometric properties of the Cognitive Flexibility Inventory by controlling for these plausible response style biases combining positive and negative items for both the Alternatives and Control subscales (i.e., rephrasing half of the items in both scales).

Encouragingly, our analyses showed evidence of reliability for both subscales, with reliability coefficients across dimensions above .86. These findings are similar to those obtained by other adaptations of the Cognitive Flexibility Inventory, which support the reliability of the internal consistency of the scale (Gülüm & Dağ, 2012; Kurginyan & Osavolyuk, 2018; Oshiro et al., 2016; Portoghese et al., 2020). Another interesting finding was that we did not find differences in CF between men and women, which is consistent with previous research (M. M. Martin & Rubin, 1995). In addition, retest analyses also showed evidence about the internal stability of the Cognitive Flexibility Inventory, thus supporting the idea that CF – as evaluated with this inventory – is likely to be relatively stable over time. This is important for clinical purposes, as it suggests that cognitive inflexibility should be a therapeutic target as it might remain relatively stable unless addressed with psychological treatment.

The findings of this study also provided evidence of the construct (convergent/divergent) validity of the Cognitive Flexibility Inventory in relation to other psychological constructs (emotion regulation, attributional style, and depressive symptoms). Previous validation studies of the Cognitive Flexibility Inventory also evidenced that CF is positively associated with the use of adaptive coping styles and negatively related to the severity of depression symptomatology (Dennis & Vander Wal, 2010; Kurginyan & Osavolyuk, 2018). The present study and previous results are consistent with the idea that

**Table 4.** Construct Validity Evidences of the Cognitive Flexibility Inventory in Relation to Other Measures

	1	2	3	4	5	6	7	8	9	10	11	12	13
1. CFI Alternatives	1												
2. CFI Control	.15*	1											
3. CERQ-A	.10	-.08	1										
4. CERQ-PR	.21***	.16**	.16**	1									
5. CERQ-PRC	.34***	.20***	.26***	.50***	1								
6. CERQ-PP	.24***	.09**	.30***	.36***	.56***	1							
7. CERQ-SB	.072	-.20***	.14**	-.10	-.01	.20	1						
8. CERQ-R	.19***	-.07	.16**	-.01	.13*	-.01	.31***	1					
9. CERQ-C	-.20***	-.30**	.11*	-.13*	-.21***	-.22***	.32***	.27**	1				
10. CERQ-OB	-.09	-.01	-.02	-.04	-.07	-.05	-.08	.13**	.25***	1			
11. CERQ-RP	.49***	.19***	.22	.34***	.59***	.36***	.04	.22***	-.18***	.01	1		
12. ASQ Positive style	.28***	.20**	-.01	.26***	.31***	.04	-.10	.19**	-.09	-.03	.20**	1	
13. ASQ Negative Style	-.14*	-.19**	-.03	-.19**	-.15*	-.07	.39***	.12	.30***	.04	-.19**	.09	1
14. BDI-II Depression	-.22**	-.34***	-.05	-.32***	-.39***	-.22***	.39***	.24**	.52**	.07**	-.42**	-.23***	.40***

Note. CFI = Cognitive Flexibility Inventory; CERQ-A = Acceptance, CERQ-PRC = Positive reappraisal, CERQ-PR = Positive refocusing, CERQ-PP = Putting into perspective, CERQ-C Catastrophizing; CERQ-SB Self-Blame, CERQ-OB Other-blame, CERQ-R Rumination, and CERQ-RP Refocusing on Planning; ASQ Attributional Styles Questionnaire; BDI-II Beck Depression Inventory-II.

\* $p < .05$ , \*\* $p < .01$ , \*\*\* $p < .001$ .

cognitive inflexibility might be a transdiagnostic process associated with the use of certain emotion regulation strategies and a better emotional adaptation to face difficulties, which is in line with the theoretical model of psychopathology proposed by the Unified Protocol for the transdiagnostic treatment of emotional disorders (Boisseau et al., 2010).

Interestingly, an unexpected positive association was found between CF and rumination. Specifically, higher cognitive flexibility was associated with higher rumination, which has been considered a maladaptive strategy of emotion regulation (Aldao et al., 2010). Rumination can be defined as a repetitive thinking about the reasons and consequences of one's problems without taking action (Nolen-Hoeksema et al., 2008). This construct consists of two components: brooding and reflection. The former increases depressive feelings by comparison of one's current situation with some unachieved standard, whereas the latter appears to be related to engaging in cognitive problem solving to alleviate one's depressive symptoms (Bastin et al., 2014; Treynor et al., 2003). Therefore, and taking into account that our sample is composed of healthy individuals, engaging in adaptive reflection for responding to contextual demands could be adaptive (L. L. Martin & Tesser, 1996), and it could help understand the positive association between CF and rumination. In any case, this interesting finding should be better explored in further research.

Regarding the association between the Cognitive Flexibility Inventory subscales and the measure of attributional styles, we found that the Cognitive Flexibility Inventory was directly associated with the positive-composite score of the ASQ and inversely associated with its negative-composite scale. Therefore, we found that individuals with higher CF tend to perceive that the causes of positive events are stable, controllable, and dependent on themselves, whereas high scores in the Cognitive Flexibility Inventory were associated with a view of negative events as more dependent on causes that are external, unstable, and context-specific. Past research combined positive and negative attributional scores to obtain a total measure of CF (Dennis & Vander Wal, 2010; Kurginyan & Osavolyuk, 2018; Teasdale et al., 2001). However, the ASQ was not developed to be a measure of CF, so the total score may not be appropriate and suitable for this purpose (Fresco et al., 2007). This might explain why past research has obtained unexpected findings when exploring the relationship between the scores obtained with the Cognitive Flexibility Inventory and the ASQ (Dennis & Vander Wal, 2010; Kurginyan & Osavolyuk, 2018), which was interpreted as evidence that CF in the Cognitive Flexibility Inventory was associated with more inflexibility in the ASQ. Based on the results of this study, it is not recommended to use a global computation based on extreme scores on the ASQ as a measure of CF, as it may lead to unexpected findings.

The present study might have made important contributions to the literature on CF. However, a number of limitations should also be considered. For instance, while this instrument might be especially useful for clinical practice, the validation was conducted among college students. Previous validations of the Cognitive Flexibility Inventory have also been conducted with non-clinical samples (Dennis & Vander Wal, 2010; Gülüm & Dağ, 2012; Kurginyan & Osavolyuk, 2018; Oshiro et al., 2016; Portoghese et al., 2020; Shareh et al., 2014), which might be a good starting point for further research. However, it would be recommended to test the validity and reliability of this measure in a clinical sample, including individuals suffering from affective disorders characterized by rigid thinking (e.g., depression, generalized anxiety disorder, or obsessive-compulsive disorder), and particularly to compare CF in clinical and non-clinical populations. In addition, further studies are needed to confirm the current structure of the Spanish version of the Cognitive Flexibility Inventory using confirmatory factor analyses in larger samples.

Despite these limitations, the results of the present study might be important for research and clinical purposes. In particular,

this study showed evidence of the validity and reliability of the Spanish adaptation of the Cognitive Flexibility Inventory, enabling their use in clinical and research contexts. This might be an important contribution for professionals interested in a wide range of different disorders, as this instrument could help assess the ability to cognitively adapt and face adverse events, which is a core trait to predict greater therapeutic success in a wide range of psychotherapeutic approaches.

### Conflict of Interest

The authors of this article declare no conflict of interest.

### Note

The datasets generated during and/or analysed during the current study are available at [https://osf.io/3fbx5/?view\\_only=119640dc32bf47b2b508718f180cf165](https://osf.io/3fbx5/?view_only=119640dc32bf47b2b508718f180cf165)

### References

- Aldao, A., Nolen-Hoeksema, S., & Schweizer, S. (2010). Emotion-regulation strategies across psychopathology: A meta-analytic review. *Clinical Psychology Review, 30*(2), 217-237. <https://doi.org/10.1016/j.cpr.2009.11.004>
- Allport, D. A., Styles, E. A., & Hsieh, S. (1994). Shifting intentional set: Exploring the dynamic control of tasks. In C. Umiltà & M. Moscovitch (Eds.), *Attention and performance 15: Conscious and nonconscious information processing* (pp. 421-452). The MIT Press.
- Bastin, M., Bijttebier, P., Raes, F., & Vasey, M. W. (2014). Brooding and reflecting in an interpersonal context. *Personality and Individual Differences, 63*, 100-105. <https://doi.org/10.1016/j.paid.2014.01.062>
- Beck, A.T., Rush, A.J., Shaw, B.F. & Emery, G. (1979) *Cognitive therapy of depression*. Guilford Press.
- Beck, A. T., Steer, R. A., & Brown, G. (1996). Beck Depression Inventory -II (BDI-II) [Database record]. APA PsycTests. <https://doi.org/10.1037/t00742-000>
- Boisseau, C. L., Farchione, T. J., Fairholme, C. P., Ellard, K., K., & Barlow, D. (2010). The development of the unified protocol for the transdiagnostic treatment of emotional disorders: A case study. *Cognitive and Behavioral Practice, 17*(1), 102-113. <https://doi.org/10.1016/j.cbpra.2009.09.003>
- Bonanno, G. A., & Burton, C. L. (2013). Regulatory flexibility: An individual differences perspective on coping and emotion regulation. *Perspectives on Psychological Science, 8*(6), 591-612. <https://doi.org/10.1177/1745691613504116>
- Brislin, R. W. (1976). Comparative research methodology: Cross-cultural studies. *International Journal of Psychology, 11*(3), 215 -229. <https://doi.org/10.1080/00207597608247359>
- Brown, T. A. (2003). Confirmatory factor analysis of the Penn State Worry Questionnaire: Multiple factors or method effects? *Behaviour Research and Therapy, 41*(12), 1411-1426. [https://doi.org/10.1016/S0005-7967\(03\)00059-7](https://doi.org/10.1016/S0005-7967(03)00059-7)
- Checa, I., & Espejo, B. (2018). Method effects associated with reversed items in the 29 Items Spanish version of Ryff's Well-Being Scales. *Neuropsychiatry, 8*(5), 1533-1540. <https://doi.org/10.4172/neuropsychiatry.1000488>
- Cohen, J. (1992). A power primer. *Psychological Bulletin, 112*(1), 155-159. <https://doi.org/10.1037/0033-2909.112.1.155>
- Deák, G. O. (2003). The development of cognitive flexibility and language abilities. In R. V. Kail (Ed.), *Advances in child development and behavior* (pp. 271-327). Academic Press.
- Demirtaş, A. S. (2020). Cognitive flexibility and mental well-being in Turkish adolescents: The mediating role of academic, social, and emotional self-efficacy. *Anales de Psicología, 36*(1), 111-121. <https://dx.doi.org/10.6018/analesps.36.1.336681>
- Dennis, J. P., & Vander Wal, J. S. (2010). The cognitive flexibility inventory: Instrument development and estimates of reliability and validity. *Cognitive Therapy and Research, 34*(3), 241-253. <https://doi.org/10.1007/s10608-009-9276-4>
- Domínguez-Sánchez, F. J., Lasa-Aristu, A., Amor, P. J., & Holgado-Tello, F. P. (2013). Psychometric properties of the Spanish version of the Cognitive Emotion Regulation Questionnaire. *Assessment, 20*(2), 253-261. <https://doi.org/10.1177/1073191110397274>
- Fresco, D. M., Rytwinski, N. K., & Craighead, L. W. (2007). Explanatory flexibility and negative life events interact to predict depression symptoms. *Journal of Social and Clinical Psychology, 26*(5), 595-608. <https://doi.org/10.1521/jscp.2007.26.5.595>
- García-García, M., Barceló, F., Clemente, I. C., & Escera, C. (2010). The role of the dopamine transporter DAT1 genotype on the neural correlates of

- cognitive flexibility. *European Journal of Neuroscience*, 31(4), 754-760. <https://doi.org/10.1111/j.1460-9568.2010.07102.x>
- Garnefski, N., Kraaij, V., & Spinhoven, P. (2001). Negative life events, cognitive emotion regulation and emotional problems. *Personality and Individual Differences*, 30(8), 1311-1327. [https://doi.org/10.1016/S0191-8869\(00\)00113-6](https://doi.org/10.1016/S0191-8869(00)00113-6)
- Geurts, H. M., Corbett, B., & Solomon, M. (2009). The paradox of cognitive flexibility in autism. *Trends in Cognitive Sciences*, 13(2), 74-82. <https://doi.org/10.1016/j.tics.2008.11.006>
- Guadagnoli, E., & Velicer, W. F. (1988). Relation of sample size to the stability of component patterns. *Psychological Bulletin*, 103(2), 265-275. <https://doi.org/10.1037/0033-2909.103.2.265>
- Gülüm, İ. V., & Dağ, İ. (2012). The Turkish adaptation, validity and reliability study of the Repetitive Thinking Questionnaire and the Cognitive Flexibility Inventory. *Anatolian Journal of Psychiatry*, 13, 216-223.
- Ionescu, T. (2012). Exploring the nature of cognitive flexibility. *New Ideas in Psychology*, 30(2), 190-200. <https://doi.org/10.1016/j.newideapsych.2011.11.001>
- Ionescu, T. (2017). The variability-stability-flexibility pattern: A possible key to understanding the flexibility of the human mind. *Review of General Psychology*, 21(2), 123-131. <https://doi.org/10.1037/gpr0000110>
- Jones R. N. (2006). Identification of measurement differences between English and Spanish language versions of the Mini-Mental State Examination. Detecting differential item functioning using MIMIC modeling. *Medical care*, 44(11 Suppl 3), S124-S133. <https://doi.org/10.1097/01.mlr.0000245250.50114.0f>
- Kashdan, T. B., & Rottenberg, J. (2010). Psychological flexibility as a fundamental aspect of health. *Clinical Psychology Review*, 30(7), 865-878. <https://doi.org/10.1016/j.cpr.2010.03.001>
- Kurginyan, S. S., & Osavolyuk, E. Y. (2018). Psychometric properties of a Russian version of the Cognitive Flexibility Inventory (CFI-R). *Frontiers in Psychology*, 9, 1-9. <https://doi.org/10.3389/fpsyg.2018.00845>
- Marsh, H. W., Hau, K.-T., & Wen, Z. (2004). In search of golden rules: Comment on hypothesis-testing approaches to setting cutoff values for fit indexes and dangers in overgeneralizing Hu and Bentler's (1999) findings. *Structural Equation Modeling: A Multidisciplinary Journal*, 11(3), 320-341. [https://doi.org/10.1207/s15328007sem1103\\_2](https://doi.org/10.1207/s15328007sem1103_2)
- Martin, L. L., & Tesser, A. (1996). Some ruminative thoughts. In R. S. Wyer (Ed.), *Ruminative thoughts* (pp. 1-47). Lawrence Erlbaum Associates, Inc.
- Martin, M. M., & Anderson, C. M. (1998). The cognitive flexibility scale: Three validity studies. *International Journal of Phytoremediation*, 21(1), 1-9. <https://doi.org/10.1080/08934219809367680>
- Martin, M. M., & Rubin, R. B. (1995). A new measure of cognitive flexibility. *Psychological Reports*, 76(2), 623-626. <https://doi.org/10.2466/pr0.1995.76.2.623>
- McCaffrey, R. J., Ortega, A., Orsillo, S. M., Nelles, W. B., & Haase, R. F. (1992). Practice effects in repeated neuropsychological assessments. *Clinical Neuropsychologist*, 6(1), 32-42. <https://doi.org/10.1080/13854049208404115>
- Miyake, A., Friedman, N. P., Emerson, M. J., Witzki, A. H., Howerter, A., & Wager, T. D. (2000). The unity and diversity of executive functions and their contributions to complex "Frontal Lobe" tasks: A latent variable analysis. *Cognitive Psychology*, 41(1), 49-100. <https://doi.org/10.1006/cogp.1999.0734>
- Nolen-Hoeksema, S., Wisco, B. E., & Lyubomirsky, S. (2008). Rethinking rumination. *Perspectives on Psychological Science: A Journal of the Association for Psychological Science*, 3(5), 400-424. <https://doi.org/10.1111/j.1745-6924.2008.00088.x>
- Oshiro, K., Nagaoka, S., & Shimizu, E. (2016). Development and validation of the Japanese version of cognitive flexibility scale. *BMC Research Notes*, 9(1), 1-8. <https://doi.org/10.1186/s13104-016-2070-y>
- Peterson, C., Semmel, A., von Baeyer, C., Abramson, L. Y., Metalsky, G. I., & Seligman, M. E. P. (1982). The Attributional Style Questionnaire. *Cognitive Therapy and Research*, 6(3), 287-299. <https://doi.org/10.1007/BF01173577>
- Portoghese, I., Lasio, M., Conti, R., Mascia, M. L., Hitchcott, P., Agus, M., Gemignani, A., & Penna, M. P. (2020). Cognitive Flexibility Inventory: Factor structure, invariance, reliability, convergent, and discriminant validity among Italian university students. *PsyCh Journal*, 9(6), 934-941. <https://doi.org/10.1002/pchj.401>
- Rogers, R. D., & Monsell, S. (1995). Costs of a predictable switch between simple cognitive tasks. *Journal of Experimental Psychology: General*, 124(2), 207-231. <https://doi.org/10.1037/0096-3445.124.2.207>
- Sanjuán, P. & Magallares, A. (2008, July). *Reliability and factorial validity of Attributional Style Questionnaire in a Spanish sample*. VIII European Conference of Psychological Assessment, Budapest, Hungary.
- Sanz, J., Perdigón, A. L., & Vázquez, C. (2003). Adaptación española del Inventario para la Depresión de Beck-II (BDI-II): 2. Propiedades psicométricas en población general. *Clínica y Salud*, 14(3), 249-280.
- Shareh, H., Farmani, A., Soltani, & E. (2014). Investigating the reliability and validity of the Cognitive Flexibility Inventory (CFI-I) among Iranian university students. *Practice in Clinical Psychology*, 2(1), 43-50. <https://doi.org/10.5430/ijelt.v1n2p53>
- Spector, P. E., Katwyk, P. T. Van, Brannick, M. T., & Chen, P. Y. (1997). When two factors don't reflect two constructs: How item characteristics can produce artifactual factors. *Journal of Management*, 23(5), 659-677. <https://doi.org/10.1177/014920639702300503>
- Suárez-Alvarez, J., Pedrosa, I., Lozano, L. M., García-Cueto, E., Cuesta, M., & Muñiz, J. (2018). Using reversed items in likert scales: A questionable practice. *Psicothema*, 30(2), 149-158. <https://doi.org/10.7334/psicothema2018.33>
- Teasdale, J. D., Scott, J., Moore, R. G., Hayhurst, H., Pope, M., & Paykel, E. S. (2001). How does cognitive therapy prevent relapse in residual depression? Evidence from a controlled trial. *Journal of Consulting and Clinical Psychology*, 69(3), 347-357. <https://doi.org/10.1037/0022-006x.69.3.347>
- Treynor, W., Gonzalez, R., & Nolen-Hoeksema, S. (2003). Ruminative reconsider: A psychometric analysis. *Cognitive Therapy and Research*, 27(3), 247-259. <https://doi.org/10.1023/A:1023910315561>
- Wang, Y., Yang, Y., Xiao, W.-T., & Su, Q. (2016). Validity and reliability of the Chinese version of the Cognitive Flexibility Inventory in college students. *Chinese Mental Health Journal*, 30(1), 58-62.
- Waugh, C. E., Thompson, R. J., & Gotlib, I. H. (2011). Flexible emotional responsiveness in trait resilience. *Emotion*, 11(5), 1059-1067. <https://doi.org/10.1037/a0021786>
- Wersebe, H., Lieb, R., Meyer, A. H., Hofer, P., & Gloster, A. T. (2018). The link between stress, well-being, and psychological flexibility during an acceptance and commitment therapy self-help intervention. *International Journal of Clinical and Health Psychology*, 18(1), 60-68. <https://doi.org/10.1016/j.ijchp.2017.09.002>
- Weston, R., & Gore, P. A., Jr. (2006). A brief guide to structural equation modeling. *The Counseling Psychologist*, 34(5), 719-751. <https://doi.org/10.1177/0011000006286345>

**Appendix A****Cognitive Flexibility Inventory (CFI)**

Por favor, utilice la siguiente escala para indicar en qué medida está de acuerdo o desacuerdo con las siguientes afirmaciones

1	2	3	4	5	6	7
<i>Totalmente en desacuerdo</i>	<i>En desacuerdo</i>	<i>Algo en desacuerdo</i>	<i>Ni acuerdo ni en desacuerdo</i>	<i>Algo de acuerdo</i>	<i>De acuerdo</i>	<i>Totalmente de acuerdo</i>
1. Soy bueno evaluando situaciones						
2. Tengo dificultades para tomar decisiones cuando me enfrento a situaciones difíciles						
3. Considero múltiples opciones cuando tomo una decisión						
4. Cuando me enfrento a situaciones difíciles, me siento como si perdiera el control						
5. Me gusta ver las situaciones difíciles desde muchos ángulos diferentes						
6. Busco información adicional que no está disponible inmediatamente antes de atribuir causas al comportamiento						
7. Cuando me encuentro ante situaciones difíciles, me estreso tanto que no puedo pensar en una manera de resolver la situación						
8. Intento pensar sobre las cosas desde el punto de vista de otra persona.						
9. Encuentro problemático que haya tantas formas de lidiar con situaciones difíciles						
10. Soy bueno poniéndome en la piel de otra gente						
11. Cuando me enfrento a situaciones difíciles, no sé qué hacer.						
12. Es importante ver las situaciones desde muchos ángulos.						
13. Cuando estoy en situaciones difíciles, considero múltiples opciones antes de decidir cómo comportarme.						
14. A menudo veo una situación desde diferentes puntos de vista.						
15. Soy capaz de superar las dificultades a las que me enfrento en la vida						
16. Considero todos los hechos e información disponibles cuando atribuyo causas al comportamiento.						
17. Siento que no tengo poder para cambiar las cosas en situaciones difíciles						
18. Cuando me enfrento con situaciones difíciles, paro e intento pensar en diferentes formas de resolverlas.						
19. Puedo pensar en más de una manera de resolver una situación difícil cuando me enfrento a ella.						
20. Considero múltiples opciones cuando respondo a una situación difícil.						



## Appendix B

## Model Estimated Covariances/Correlations/Residual Correlations

	CFL_1	CFL_2R	CFL_3	CFL_4R	CFL_5
CFL_1					
CFL_2R	0.154				
CFL_3	0.310	-0.063			
CFL_4R	0.198	0.614	-0.029		
CFL_5	0.342	0.009	0.442	0.051	
CFL_6	0.308	0.132	0.340	0.176	0.371
CFL_7R	0.206	0.584	-0.005	0.641	0.072
CFL_8	0.208	-0.043	0.291	-0.021	0.297
CFL_9R	0.118	0.543	-0.100	0.588	-0.033
CFL_10	0.203	-0.036	0.281	-0.013	0.288
CFL_11R	0.230	0.680	-0.019	0.745	0.070
CFL_12	0.310	0.055	0.378	0.095	0.401
CFL_13	0.403	-0.004	0.528	0.044	0.549
CFL_14	0.425	-0.013	0.559	0.038	0.580
CFL_15	0.270	0.105	0.303	0.143	0.328
CFL_16	0.341	0.094	0.401	0.140	0.429
CFL_17R	0.149	0.503	-0.042	0.549	0.023
CFL_18	0.425	0.101	0.507	0.158	0.540
CFL_19	0.417	0.129	0.484	0.186	0.520
CFL_20	0.454	0.084	0.553	0.143	0.586
	CFL_6	CFL_7R	CFL_8	CFL_9R	CFL_10
CFL_7R	0.187				
CFL_8	0.228	-0.004			
CFL_9R	0.094	0.557	-0.068		
CFL_10	0.222	0.002	0.189	-0.060	
CFL_11R	0.207	0.712	-0.014	0.650	-0.006
CFL_12	0.333	0.111	0.254	0.017	0.247
CFL_13	0.438	0.070	0.355	-0.053	0.344
CFL_14	0.461	0.066	0.376	-0.064	0.364
CFL_15	0.286	0.153	0.203	0.072	0.198
CFL_16	0.364	0.155	0.269	0.052	0.262
CFL_17R	0.130	0.523	-0.029	0.484	-0.023
CFL_18	0.454	0.178	0.340	0.050	0.331
CFL_19	0.444	0.204	0.325	0.078	0.316
CFL_20	0.487	0.166	0.371	0.029	0.360
	CFL_11R	CFL_12	CFL_13	CFL_14	CFL_15
CFL_12	0.118				
CFL_13	0.065	0.476			
CFL_14	0.059	0.502	0.691		
CFL_15	0.169	0.294	0.388	0.409	
CFL_16	0.168	0.380	0.508	0.536	0.320
CFL_17R	0.609	0.062	0.014	0.008	0.104
CFL_18	0.191	0.477	0.640	0.676	0.401
CFL_19	0.222	0.462	0.615	0.649	0.391
CFL_20	0.176	0.515	0.695	0.734	0.430
	CFL_16	CFL_17R	CFL_18	CFL_19	CFL_20
CFL_17R	0.097				
CFL_18	0.516	0.108			
CFL_19	0.502	0.132	0.628		
CFL_20	0.556	0.094	0.698	0.676	

## Model Estimated Covariances/Correlations/Residual Correlations

	CFI_1	CFI_2R	CFI_3	CFI_4R	CFI_5
CFI_1					
CFI_2R	0.012				
CFI_3	0.010	-0.167			
CFI_4R	-0.006	0.013	-0.020		
CFI_5	0.040	0.026	0.052	-0.071	
CFI_6	-0.055	0.029	-0.108	-0.055	0.108
CFI_7R	0.011	-0.031	0.062	0.049	0.034
CFI_8	-0.068	0.047	-0.029	0.039	0.039
CFI_9R	-0.025	-0.008	0.100	-0.055	0.048
CFI_10	0.111	0.105	0.002	0.024	-0.059
CFI_11R	0.000	-0.002	0.081	-0.007	0.050
CFI_12	0.086	0.067	-0.055	-0.025	0.093
CFI_13	-0.142	-0.029	0.046	0.069	-0.020
CFI_14	-0.014	0.025	-0.011	0.021	0.072
CFI_15	0.168	0.016	-0.014	-0.034	-0.060
CFI_16	0.013	0.055	-0.109	-0.042	-0.018
CFI_17R	-0.045	0.009	0.061	-0.022	0.002
CFI_18	0.055	0.039	-0.008	0.000	-0.040
CFI_19	-0.027	0.022	-0.020	0.004	-0.072
CFI_20	-0.047	-0.070	0.057	0.035	-0.077
	CFI_6	CFI_7R	CFI_8	CFI_9R	CFI_10
CFI_7R	-0.008				
CFI_8	0.037	-0.127			
CFI_9R	0.014	0.002	-0.024		
CFI_10	-0.004	-0.061	0.118	-0.163	
CFI_11R	0.006	-0.028	0.086	0.040	0.036
CFI_12	0.064	-0.013	0.071	0.002	0.070
CFI_13	0.005	0.022	-0.057	0.051	-0.033
CFI_14	0.018	0.022	0.105	0.000	0.016
CFI_15	-0.116	-0.004	-0.095	-0.041	0.092
CFI_16	0.197	-0.021	0.030	-0.038	0.046
CFI_17R	0.039	0.005	0.050	-0.017	0.122
CFI_18	-0.053	0.034	-0.054	0.030	-0.078
CFI_19	-0.088	-0.044	-0.042	-0.005	-0.059
CFI_20	-0.039	0.010	-0.067	0.007	-0.066
	CFI_11R	CFI_12	CFI_13	CFI_14	CFI_15
CFI_12	-0.070				
CFI_13	-0.016	-0.068			
CFI_14	0.020	-0.003	0.043		
CFI_15	-0.063	0.064	-0.196	-0.076	
CFI_16	-0.057	0.017	0.008	-0.024	0.008
CFI_17R	0.025	0.060	-0.009	0.020	0.022
CFI_18	0.005	-0.041	-0.004	-0.034	0.107
CFI_19	-0.015	-0.035	-0.003	-0.077	0.052
CFI_20	-0.008	-0.058	0.037	-0.023	-0.034
	CFI_16	CFI_17R	CFI_18	CFI_19	CFI_20
CFI_17R	0.029				
CFI_18	-0.042	-0.121			
CFI_19	-0.043	-0.020	0.057		
CFI_20	-0.021	0.027	0.000	0.072	