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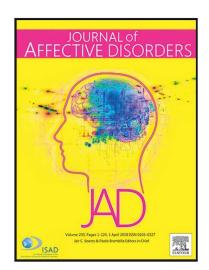
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Highlights

- This is the first study to explore the psychometric properties of the OASIS and ODSIS, screening tools for anxiety and mood disorders, in a Spanish sample of patients diagnosed with emotional disorders in a specialized mental health setting.
- The OASIS and ODSIS are easily administered screening tools with an excellent internal
 consistency, favorable convergent and discriminant validity, and good sensitivity to
 therapeutic change.
- A cut-off point of 10 for both instruments results in the best rate of correct clinical/nonclinical classifications.
- The Spanish adaptations of the OASIS and the ODSIS demonstrated their cross-cultural utility.
- The OASIS and ODSIS should make screening of anxiety and depressive symptoms in public health settings more feasible and could facilitate adequate referrals and routine evaluation.



Psychometric properties and validation of the Spanish versions of the Overall Anxiety and Depression Severity and Impairment Scales

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Abstract

Background: Anxiety and depressive disorders are the most frequent disorders for which patients seek care in public health settings in Spain. This study aimed at validating the Overall Anxiety Severity and Impairment Scale (OASIS) and the Overall Depression Severity and Impairment Scale (ODSIS), which are brief screening scales for anxiety and depression consisting of only five items each.

Methods: The study was conducted in a Spanish clinical sample receiving outpatient mental health treatment (N = 339). A subsample of participants (n = 219) was assessed before and after receiving a course of cognitive-behavioral treatment.

Results: The results revealed excellent internal consistency estimates (Cronbach's alpha for the OASIS and the ODSIS was .87 and .94, respectively), along with promising convergent and discriminant validity and test-criterion relationships (i.e., moderate correlation with other measures of depression and anxiety, as well as with neuroticism, quality of life, adjustment, and negative affect). A one-dimensional structure was obtained for the OASIS and the ODSIS. The ROC analyses indicated an area under the curve of .83 for the OASIS and the ODSIS when predicting moderate-to-severe anxiety and depression, respectively. Good sensitivity to therapeutic change was also evidence and the analysis of the sensitivity as a function of 1-specificity area suggested a cutoff value of 10 for both scales.

Limitations: Inter-rater reliability of diagnoses with the ADIS-IV interview could not be investigated and the results obtained may not be generalizable to other samples and health settings.

Conclusions: The availability of these two short and psychometrically sound measures should make screening of anxiety and depressive symptoms in routine care more feasible.

Key words: anxiety; depression; transdiagnostic; screening; validation; psychometric properties.

Introduction

Anxiety and depressive disorders, also known as emotional disorders (EDs; Barlow, 1991) are common conditions that negatively impact functioning (Campbell-Sills et al., 2009). According to the World Health Organization (WHO), in Spain alone, anxiety and depressive disorders affect nearly 2 million and 2.5 million people, respectively, which amount to up to 5.2% of the population in the country (WHO, 2017). Not surprisingly, EDs are largely prevalent in primary care and specialized mental health settings in Spain (Gutiérrez-Fraile et al., 2011; Muñoz-Navarro et al., 2017; Ruiz-Rodríguez et al., 2017). The present study aims at adapting two short screening tools for anxiety and depressive symptoms to be used in Spanish public care.

Several authors have indicated that diagnosis and monitoring of patients with EDs in Spanish public health settings is problematic. For example, a study by Castro-Rodríguez and colleagues (2015) suggested that half of the patients with an ED are not correctly diagnosed in primary care. Consequently, a large number of individuals with ED are not properly referred to specialized care and do not receive the most optimal treatment (Martín Jurado et al., 2012). While there might be several factors explaining this, some authors have attributed the previous to healthcare overload, time

constraints during consultations, and family physicians' insufficient familiarity with the assessment and treatment of psychiatric disorders (Muñoz-Navarro et al., 2017). —It is important to note that it is at this first level of care where correct identification of those who need specialized interventions is of paramount importance, so quick and straightforward assessment tools for an early detection of EDs are fundamental (Cano Vindel et al., 2012). A similar problem occurs in specialized mental health settings in Spain, where limited resources (e.g., reduced number of professionals, especially psychologists) has shown to negatively impact the quality of services provided (Gabilondo et al., 2011). An example of this is the long waiting lists that currently exist in Spanish public mental health settings, that is, from 45 to up to 75 days between appointments (Chueca et al., 2003; González et al., 2008). A quick and feasible assessment tool for the evaluation of anxiety and depressive symptoms to be used in routine care would benefit psychologists and psychiatrists by reducing the time they require for the assessment of the patients' progress through treatment.

Many self-report screening measures have been designed to assess specific diagnostic categories (e.g., Panic Disorder Severity Scale [PDSS]; Shear et al., 2001). Although these scales are useful to ascertain whether an individual is likely to meet criteria for a specific disorder, the administration of a large number of screening measures (e.g., one for each ED diagnosis) is not feasible in public health settings. Thus, transdiagnostic measures that capture the spectrum of ED presentations may be particularly useful in such settings (Campbell-Sills et al., 2009). Furthermore, traditional self-report measures typically assume that the frequency of symptoms reflects the severity of the disorder, while ignoring the key role of functional impairment or distress for diagnosis (Ito et al., 2014).

The Overall Anxiety Severity and Impairment Scale (OASIS; Norman et al., 2006) and the Overall Depression Severity and Impairment Scale (ODSIS; Bentley et al., 2014) were developed in response to the limitations of existing screening tools. Both instruments were designed to assess transdiagnostic symptoms and functioning associated with anxiety and depressive disorders, respectively. The OASIS and the ODSIS have five items each and can be administered in only 2-3 minutes. Both measures have been shown to be useful in the assessment of patients' progress through treatment (Barlow et al., 2017; Norman et al., 2013; Osma et al., 2015). The OASIS has been validated into English (Norman et al., 2006), Japanese (Ito et al., 2014), and Dutch (Hermans et al., 2014). An online version of this measure has also been recently been validated into Spanish in a sample of individuals with milder depression and anxiety (González-Robles et al., 2018). A single-factor model including all items has been replicated across investigations (Bragdon et al., 2016; Campbell-Sills et al., 2009; González-Robles et al., 2018; Hermans et al., 2014; Ito et al., 2014; Moore et al., 2015; Norman et al., 2013; Norman et al., 2011; Norman et al., 2014; Ito et al., 2014; Norman et al., 2006).

The cut-off points for the OASIS have ranged from 5 to 10 in clinical samples (Bragdon et al., 2016; Campbell-Sills et al., 2009; González-Robles et al., 2018; Hermans et al., 2014; Moore et al., 2015; Norman et al., 2013) and from 8 to 9 in studies with clinical and non-clinical populations altogether (Ito et al, 2014; Norman et al., 2011).

Similar to the OASIS, the ODSIS has been validated into English (Bentley et al., 2014) and Japanese (Ito et al., 2015). A single-factor solution has been proposed for the ODSIS and internal consistency findings have been excellent (Bentley et al., 2014; Ito et al., 2015). The cutoff points for the ODSIS have ranged from between 7 and 11 with clinical and non-clinical samples altogether (Bentley et al., 2014; Ito et al., 2015).

The present study goal was to replicate and extend prior research on the psychometric properties of the OASIS and the ODSIS in a Spanish outpatient sample of mental health patients with EDs. Specifically, we aimed at replicating the findings of the validations conducted in the USA, the Netherlands, Spain, and Japan, thus providing further evidence for the transcultural validity of these scales. The study also had a novel goal, that is, to determine the sensitivity of both measures to therapeutic change after a transdiagnostic, cognitive-behavioral treatment for EDs, namely, the Unified Protocol (Barlow et al., 2017; Blind note).

The following hypotheses were proposed: 1) for both scales, a one-dimensional structure will be replicated; 2) an acceptable internal consistency will be obtained; 3) convergent and discriminant validity will be established through significant positive correlations between the OASIS and the ODSIS and measures assessing similar constructs, as well as weaker correlations with measures of different constructs; 4) the ROC curve will reveal an optimal cut-off point of approximately 10 for both measures; and finally, 5) adequate sensitivity to therapeutic change will be indicated by significant and positive correlations between the change detected by the reference measurements (the Beck Anxiety Inventory and Beck Depression Inventory-II) and the change detected by the OASIS and the ODSIS.

Method

Participants

Participants were drawn from a large effectiveness trial of the Unified Protocol (UP) in the Spanish public mental health system (Blind note). The sample for the present study was composed of 339 participants of whom 272 were women (80.2%). Mean age in the sample was 42.63 years (SD = 12.79, range = 18-77). Table 1 shows additional socio-demographic characteristics. The selected participants adults seeking psychological assistance in the national health system. The sample was obtained in 8 public mental health centers of Spain after the approval of their respective ethics and research committees (for more information, Blind note): Blind note. All participants had a primary ED diagnosis. The inclusion criteria to participate in the study were: 1) Principal diagnosis (most interfering and severe) of anxiety disorder, mood disorder, adjustment disorder; 2) The patient is over 18 years of age; 3) The patient is fluent in the language in which therapy is performed (Spanish in the present study); 4) The patient is able to attend to the evaluation and treatment sessions and signs the informed consent form; 5) Patients taking pharmacological treatment for their emotional disorder are asked to maintain the same dosages and medications for at least 3 months prior to enrolling in the study and during the whole treatment. The exclusion criteria were: 1) The patient presents a severe condition that would require to be prioritized for treatment, so that an interaction between both interventions cannot be ruled out. These include a severe mental disorder (bipolar disorder, schizophrenia, or an organic mental disorder), suicide risk at the time of assessment, or substance use in the last three months (excluding cannabis, coffee, and / or nicotine); 2) The patient has previously received 8 or more sessions of psychological treatment with clear and identifiable CBT principles within the past 5 years.

From the total sample (N = 339), we extracted the subsample of 219 participants who completed two assessments (one before psychological treatment onset and the other at the end of treatment) to assess the sensitivity to therapeutic change of the OASIS and the ODSIS. Mean age in the sample was 43.25

years (SD = 12.21, range=18-72) and 79.5% of them were women. When comparing those who completed both evaluations against participants who only answered to one assessment, we found no significant differences in the OASIS (t = .184, p < .366), the ODSIS (t = .210, p < .834), the BAI (t = .306, p < .481), and the BDI-II (t = .901, p < .246).

Procedure

When seeking care at a Spanish public mental health center, potential participants were evaluated by professional specialists (clinical psychologists, psychiatrists, or resident psychologists) with the lifetime version of the Anxiety Disorders Interview Schedule for DSM-IV (ADIS-IV-L; Brown et al., 1994) or the Clinical Structured Interview for Axis I Disorders of the Clinical Version of the DSM-IV (SCID-I-CV; First et al., 1997). This was done to guarantee the existence of a primary diagnosis of ED. We did not use the ADIS-5 because it is not available in Spanish yet. Participants meeting the requirements for the primary study (i.e., Blind note) were given information about data confidentiality and were asked to provide written informed consent. The same day, they completed the evaluation tools in person and on paper and pencil.

The proposed treatment is a cognitive-behavioral intervention called Unified Protocol for the transdiagnostic treatment of emotional disorders (PU; Barlow et al., 2018). The UP focuses on the identification of dysfunctional emotion regulation strategies that are common to all emotional disorders and helps patients to tolerate and cope with emotions in a more adaptive way (Barlow et al., 2018). The UP consists of 8 core intervention modules that can easily be delivered in group format (Bullis et al., 2015; Osma et al., 2015). For the present study, a 12-session group format was conducted in different mental health settings in Spain. The treatment had a duration of approximately three months and patients received a 2-hour session weekly. Both the patient and the therapist manuals have been previously translated into Spanish (Barlow et al., 2015)

Instruments

Semi-structured Interview Schedule for Anxiety Disorders Lifetime version (ADIS-IV-L; Brown et al., 1994); translated into Spanish by Botella and Ballester (1997). This semistructured interview evaluates anxiety and depressive disorders, as well as somatoform and substance use disorders according to the fourth edition of the Diagnostic and Statistical Manual of Mental Disorders (DSM-IV; APA, 1994). Test-retest reliability of this interview ranges from .67 to .86 (Brown et al., 2001). Clinical Structured Interview for Axis I Disorders of the Clinical Version of the DSM-IV (SCID-I-CV; First et al., 1997), adapted and validated into Spanish by First et al. (1999). This semi-structured interview is used to establish DSM-IV Axis I diagnoses (APA, 1994). The response scale ranges from 1 (Symptom absent or false) to 3 (True or clinically present). Test-retest reliability varies from .37 to 1, depending on studies (Rogers, 2001).

Beck Depression Inventory (BDI-II; Beck et al., 1996), validated into Spanish by Sanz et al. (2003). The BDI-II assesses depressive symptoms. It consists of 21 items, each with four different statements that reflect an increase in the degree of depression. A score of 0 indicates the absence of depressive symptoms and 3 reflects the most severe levels of depression. The scale has a 0-to-63 range. In the

Spanish validation with a clinical sample, a Cronbach alpha of .89 was obtained (Sanz et al., 2005). Cronbach's alpha for the BDI-II in the present sample is .917. The cut-off point used is 20, as suggested in the literature (Sanz et al., 2003).

Beck Anxiety Inventory (BAI; Beck & Steer, 1993), validated into Spanish by Sanz and Navarro (2003). The BAI is composed of 21 items that evaluate anxiety symptoms. Responses use a 4-point Likert scale ranging from 0 (Not at all) to 3 (Severely). The total score can range from 0 to 63. The Spanish validation study in a clinical sample revealed a Cronbach alpha of .91 (Sanz et al., 2012). Cronbach's alpha for the BAI in the present sample is .926. The cut-off point used is 16, as recommended in previous research (Sanz & Navarro, 2003).

Positive and Negative Affect Scale (PANAS; Watson et al., 1988), validated into Spanish by Sandín et al. (1999). The PANAS consists of 20 items that measure both positive and negative affect, with 10 items for each dimension. Respondents indicate the extent to which they experience different emotions representing both dimensions. Items use a 5-point Likert response scale ranging from 1 (Not at all or very slightly) to 5 (Extremely). In its Spanish validation, Cronbach alphas between .87 and .91 were obtained for both men and women in positive and negative affect (Sandín et al., 1999). In the present study, Cronbach's alpha for the PANAS are .889 for Positive Affect and .893 the Negative Affect.

Quality of Life Index (QLI; Ferrans & Powers, 1985), validated in Spanish by Mezzich et al. (2000). This measure is comprised of 10 items, each of which refers to a different dimension of quality of life. Each item is evaluated on a 10-point scale ranging from 1 (Bad) to 10 (Excellent). An overall index is calculated by summing all items (total range is 10 to 100). In the Spanish validation study (Mezzich et al., 2000), the QLI obtained a Cronbach alpha of .89. The Cronbach's alpha in the present sample is .869.

Scale of Maladjustment (SM; Echeburúa et al., 2000). This scale reflects the extent to which people's current problems affect different areas of their daily lives, including work or studies, social life, leisure time, couple relationship, and day-to-day maladjustment. All items are evaluated on a 6-point response scale ranging from 0 (Not at all) to 5 (Very severe). The total score ranges from 0 to 30, and is obtained by summing all items. Higher scores indicate poorer adjustment. The scale had excellent test-retest reliability (r = .86) and an internal consistency of .94 (Echeburúa et al., 2000). The Cronbach's alpha in the present sample is .830.

NEO-Five-Factor Personality Inventory (NEO-FFI; Costa & McCrae, 1999). The NEOFFI consists of 60 items that provide a quick and general measurement of the Big Five personality dimensions. These are neuroticism, extraversion, openness to experience, agreeableness, and conscientiousness. The responses are rated on a 5-point Likert scale ranging from 0 (Strongly disagree) to 4 (Strongly agree). Each personality dimension is composed of 12 items (scores for each dimension range from 0 to 48). The internal consistency of the NEO-FFI dimensions ranges between .82 and .90. Cronbach's alpha values in the present sample are .760 for Neuroticism, .809 for Extraversion, .769 for Opening to the Experience, .709 for Agreeableness, and .708 for Conscientiousness.

Overall Anxiety Severity and Impairment Scale (OASIS; Norman et al., 2006). The scale presents five items. This items are related with the frequency of symptoms (During the last week, did you often feel anxious?), their intensity (During the last week, when you felt anxious, to what extent was your anxiety intense or severe?), their interference with the person's work or school life (During the last week, to what extent did anxiety interfere with your ability to do the things you had to do regarding work, school, or your home?), and their interference with social life (During the last week, to what extent has anxiety interfered with your social life and your relationships?). All items are rated on a 5-point Likert scale ranging from 0 (*I didn't feel anxious*) to 4 (*Constant anxiety*). The total scale score, which ranges from 0 to 20, is obtained by summing all items. Higher scores are indicative of greater severity and functional impairment as a result of anxious symptoms (Norman et al., 2006). For this study, we used the Spanish translation of the original scales (Appendix 1), which was carried out by Osma and García-Palacios by means of a back translation process. The Spanish version of the OASIS was obtained after a backward translation to ensure conceptual equivalency. Forward translation into Spanish was carried out by an independent native-speaking translator who was proficient in English. The Spanish version was back translated into English by an independent English native translator. No significant changes were required to obtain the final version. Cronbach's alpha for the OASIS in the present sample is presented in the results section.

Overall Depression Severity and Impairment Scale (ODSIS; Bentley et al., 2014). Similar to the OASIS, the ODSIS includes five items, which are related with the frequency of symptoms (During the last week, did vou often feel depressed?), their intensity (During the last week, when you felt depressed, to what extent was your depression intense or severe?), their interference with the person's work or school life (During the last week, to what extent did depression interfere with your ability to do the things you had to do regarding work, school, or your home?), and their interference with social life (During the last week, to what extent has depression interfered with your social life and your relationships?). The third item of the OASIS and ODSIS differs, however. In the OASIS, this item evaluates the presence of avoidant behavior (During the last week, how often have you avoided situations, places, objects, or activities because of your anxiety or fear?), whereas the ODSIS measures the loss of interest or difficulty in participating in activities (During the last week, how often have you had difficulty to carry out or feel interest in activities that you normally enjoy because of your depression?). Items in the ODSIS use a 5 point Likert response scale ranging from 0 (I didn't feel depressed) to 4 (Constant depression). Again, total scores range from 0 to 20 after summing all items, with higher values revealing a more severe and functional impairment as a result of depressive symptoms (Bentley et al., 2014). As with the OASIS, we used the same process to translate the Spanish version of the ODSIS, which can be found in Appendix 1. Like with the OASIS, the Cronbach's alpha value for the ODSIS in the present sample is presented in the results section.

Data analysis

Demographic and sex differences in study variables

Analyses were carried out using the statistical package IBM SPSS Statistics version

22.0 for Windows (IBM Corp., 2013) and Mplus version 6 program (Muthen & Muthen, 1998). Firstly, sociodemographic characteristics of the total sample were analyzed (N = 339), as well as the means and standard deviations for all measures. Next, sex differences in the OASIS and the ODSIS were investigated. Because scores in men and women were not normally-distributed (Kolmogorov-Smirnov normality test), a Mann-Whitney non-parametric U-test was selected for this analysis.

Model fit, internal consistency, and convergent and discriminant validity

Next, we investigated the internal structure and the convergent and discriminant validity of the OASIS and the ODSIS. The internal structure of the OASIS and the ODSIS were calculated with a Confirmatory Factor Analysis (CFA). Because we had no missing data, we were able to use a maximum likelihood estimator with standard errors and a mean-adjusted chi-square test statistic that are robust to non-normality (MLM). The one-factor model fit was evaluated with the Satorra-Bentler robust Chi-square test, the root mean square error of approximation (RMSEA), the comparative fit index (CFI), the Tucker-Lewis index (TLI), and the standardized root mean square residual

(SRMR). RMSEA and SRMR scores below .05 indicate good model fit and values below

.08 reflect acceptable fit. CFI and TLI scores above .95 reveal an excellent fit (Hu & Bentler,

1999). The convergent and discriminant validity of the OASIS and the ODSIS was explored by calculating their Pearson correlations with well-established measures of anxiety (BAI) and depression (BDI-II), respectively. To test the hypothesis that the OASIS would be more strongly associated with the BAI and the ODSIS would correlate more with the BDI, the strength of these bivariate associations was compared using the *cocor R* package (Diedenhofen & Musch, 2015). Pearson correlations with several constructs closely related to anxiety and depressive symptoms (i.e., personality, affect, quality of life, and adjustment) were also conducted to investigate the convergent and discriminant validity of the OASIS and ODSIS. Specifically, the correlation between the aforementioned related constructs and the OASIS and the ODSIS are expected to be significant (i.e., a positive association with neuroticism, negative affect, and maladjustment, and a negative correlation with extraversion, positive affect, and quality of life or similar constructs), but smaller when compared to the correlation between the OASIS and the ODSIS with the BAI and the BDI-II (e.g., Campbell-Sills et al., 2009; Ito et al., 2015).

Optimal cut-offs and sensitivity to change

As a final step, we calculated the optimal cut-off points for the two aforementioned scales using a ROC curve analysis, as well as their sensitivity to capture therapeutic change. In relation to the former goal, sensitivity and specificity were investigated for several cut-offs, following the guidelines of Zweig and Campbell (1993). Positive and negative predictive values and likelihood ratios were also investigated. The sensitivity and specificity, as well as the positive and negative predictive values and likelihood ratios of the OASIS and the ODSIS were calculated in relation to the BAI and BDI-II, respectively. The BAI and BDI-II cut-offs used to classify individuals as severely anxious and depressed were 16 and 20, respectively (Sanz, 2014; Sanz et al., 2003). Regarding sensitivity to therapeutic change, we first calculated variables for the OASIS, the ODSIS, the BAI, and the BDI-II by subtracting the total score after treatment from the total score prior to treatment. Next, Pearson correlations between OASIS and BAI, and ODSIS and BDI-II, respectively, change scores were conducted. Also, we calculated Cohen's *d* effect size to help interpret the findings. This analysis was

carried out with the subsample of participants (n = 219) that participated in two assessment points (pre-treatment and post-treatment).

Results

As observed in Table 1, the most frequent primary diagnoses in the sample (N=339) were adjustment disorders (n=77, 22.7%), followed by major depressive disorders (n=62, 18.3%), and generalized anxiety disorders (n=43, 12.7%). Ninety nine participants (26.8%) also had a secondary diagnosis, of which major depressive disorder (n=19, 19.2%) and non-specific anxiety disorder (n=14, 14.1%) were the most frequent.

Table 2 shows the mean scores and standard deviations for all instruments. The average score for the OASIS and the ODSIS was 10.45 (SD = 4.49, range 0-20) and 9.87 (SD = 5.14, range 0-20), respectively.

Model fit, internal consistency, convergent and discriminant validity, optimal cut-offs, and sensitivity to change of the OASIS

As reported in past research, the error variances of items 1 and 2 in the OASIS need to be correlated for adequate fit. This was also true in the present study. When item 1 and 2 were not correlated, a poor fit was revealed (χ^2 (5, 303) = 55.62, p< .001; RMSEA = 0.183, 90% RMSEA = 0.141-0.228; CFI = 0.926; TLI = 0.853; SRMR = 0.054). When item 1 and 2 were correlated, as indicated in previous research (Campbell-Sills et al., 2009), the fit clearly improved (χ^2 (4, 303) = 2.49, p=.646; RMSEA < 0.001, 90% RMSEA = <0.001-0.070; CFI = 1.000; TLI = 1.000; SRMR = 0.011). Factor loadings were between moderate and large and all significant at an alpha level of .001 (Table 3).

The analysis of internal consistency indicated good reliability of the OASIS ($\alpha =$

.867). Table 2 also reports the results of the convergent and discriminant validity of the OASIS, as reflected by Pearson correlations. With regard to convergent and discriminant validity, the OASIS correlated moderately both with the BAI (r = .57, p < .001), and the BDI-II (r = .60, p < .001). According to the results obtained with the $cocor\ R$ package, both correlations were comparable in strength (the 95% CI for the difference includes zero: -.0964, 0.0356). The OASIS and the ODSIS clearly correlated with each other (r = .69, p < .001). Taking the criterion validity of the OASIS, positively the scale correlated positively with a measure of neuroticism (r = .43), maladjustment (r = .57), and negative affect (r = .49), and were negatively associated with a measure of extraversion (r = .30), quality of life (r = -.57), and positive affect (r = -.34).

All correlations were significant at an alpha level of .001.

The final analyses for the OASIS included an exploration of its optimal cut-off points, followed by an analysis of its sensitivity to change. In relation to the former, the ROC analysis indicated an area under the curve of .83 when predicting moderate-to-severe anxiety as measured with the BAI (95% CI = .78, .88). We considered 10 to be the best cut-off point to discriminate between people who had moderate-to-severe anxiety symptoms and those with milder or no symptoms because of the good balance between sensitivity (.75) and specificity (.73). With this cut-off point, 74.9% of correct classifications were reached (Youden Index of .49). Positive and negative predictive values were 91.4% and 38.8%, respectively. Positive and negative likelihood ratios were 2.8 and 0.3, respectively. The analysis of

sensitivity to change (n = 219) revealed a correlation of .40 (p < .001) between the BAI and the OASIS. Cohen's d effect size was .50, showing moderate differences between assessments.

Model fit, internal consistency, convergent and discriminant validity, optimal cut-offs, and sensitivity to change of the ODSIS

Like in the OASIS, research with the ODSIS has indicated that the error variances of items 1 and 2 need to be correlated for adequate fit. We obtained the same results. Specifically, when item 1 and 2 were not correlated, a poor fit was obtained (χ 2 (5, 303) = 55.27, p< .001; RMSEA = 0.182, 90% RMSEA = 0.141-0.227; CFI = 0.957; TLI = 0.914; SRMR = 0.031). On the contrary, when item 1 and 2 were correlated, the fit clearly improved (χ 2 (4, 303) = 2.46, p= .653; RMSEA < .001, 90% RMSEA = <0.001-0.069; CFI = 1.000; TLI = 1.000; SRMR = 0.007). Factor loadings were large and significant at an alpha level of .001 (Table 3).

Similar to the OASIS, the analysis of internal consistency indicated excellent reliability of the ODSIS (α = .936). The results of the convergent and discriminant validity of the ODSIS are reported on Table 2. The ODSIS correlated moderately with the BDI-II (r = .68, p < .001) and weakly with the BAI (r = .42, p < .001). Different to the results obtained with the OASIS, the analyses indicated differences in strength between the aforementioned correlations (the 95% CI for the difference did not include zero:.1919, 0.3335). In the assessment of criterion validity, the ODSIS was found to positively correlate with a measure of neuroticism (r = .43), maladjustment (r = .58), and negative affect (r = .42), and were negatively associated with a measure of extraversion (r = -.33), quality of life (r = -.64), and positive affect (r = -.48). All correlations were significant at an alpha level of .001.

We next explored the optimal cut-off value in the ODSIS to discriminate individuals with moderate-to-severe depressive symptoms in the BDI-II from those with milder or no symptoms. The ROC analysis indicated an area under the curve of .83 (95% CI = .78, .87). Again, a cut-off of 10 was considered to provide an adequate balance between sensitivity (.74) and specificity (.79). Using this cut-off score, 75.2% of individuals were correctly classified (Youden Index of .53). Positive and negative predictive values were 91.7% and 48.4%, respectively. Positive and negative likelihood ratios were 3.5 and 0.3, respectively. Finally, the analysis of sensitivity to change (n = 219) indicated a correlation of .56 (p < .001) between the BDI-II and the ODSIS. Cohen's d effect size was .73, showing moderate differences between both evaluations.

Discussion

This is the first study to explore the psychometric properties of the OASIS and ODSIS in a Spanish sample of patients diagnosed with ED in a specialized mental health setting, and, consequently, the contributions made are innovative and necessary.

The results regarding the internal structure of the OASIS and the ODSIS, as obtained with the CFA, are consistent with other studies and indicate that a one-factor structure for the five items has an optimal fit (Bentley et al., 2014; Bragdon et al., 2016; Campbell-Sills et al., 2008; Hermans et al., 2014; Ito et al., 2014; Ito et al., 2014; Moore et al., 2015; Norman et al., 2013; Norman et al., 2011; Norman et al., 2006). The present study revealed that the influence of the covariance between Items 1 and 2 has to be taken into account when exploring the internal structure of the OASIS and the ODSIS, as noted in previous similar studies (Campbell-Sills et al., 2009). As revealed in the present

investigation, the reason for this is that the absence of anxious or depressive symptoms (score of 0 on Item 1) unequivocally leads to a score of 0 on Item 2 (intensity), so the items are dependent on each other.

Data from the present study suggests that a cut-off point of 10 for both instruments results in the best rate of correct clinical/non-clinical classifications (72.5% of correct classifications in the OASIS and 75.2% in the ODSIS), which is different to the value proposed by some previous investigations (e.g., Hermans et al., 2014; Norman et al., 2013). Both validation studies proposed cut-offs of 5 in clinical samples, with 82.5% and 91% of correct classifications respectively. As noted by these authors, the different cut-off points proposed tend to be attributable to the particular characteristics of their samples (% of women and men, variety of EDs, and % comorbidity disorders in the sample, among others) and the setting where the studies were conducted (mental health settings or community services). The previous might also explain why our recommended cut-off values are also higher than those indicated in a recent Spanish validation of the online version of the OASIS in a subclinical sample with ED, where a cut-off of 7.5 was proposed (González-Robles et al., 2018). These differences could be, at least partly, attributable to the differences in the samples and the evaluation methods used in the investigations (i.e., individuals who attended a University clinic online version of the questionnaire in the study by González-Robles, as opposed to patients who attended public mental health settings and answered to the questionnaires on a paper and pencil version in our study. In favor of this conclusion, we observed that anxiety and depressive symptoms were more severe in the present study when compared to the investigation by González-Robles and colleagues (2018). Specifically, mean scores in our sample were significantly higher in the OASIS (t = 5.78, p < .001), the ODSIS (t =6.21, p < .001), the BAI (t = 7.52, p < .001), and the BDI-II (t = 5.16, p < .001) when compared to the aforementioned investigation in Spanish settings.

Because sample characteristics are important factors in the determination of cut-off scores, we will now discuss our findings in relation to investigations including a sample of outpatients presenting a primary diagnosis of ED, similar to the present study. Our proposed cut-off point for the OASIS is the same as the one suggested by Bragdon and colleagues (2016), who with obtained a 76% of correct classifications with a cut-off of 10, and also close to the recommendation of Ito and colleagues (2015), who proposed a cut-off score of 9 with 68% of correct classifications. Somewhat lower cut-off values have also been proposed, including the work of Campbell-Sills and colleagues (2009), Moore and colleagues (2015), and Norman and colleagues (2011), who recommended a cut-off point of 8 for the OASIS, with a percentage of correct classifications of 87%, 67%, and 78%, respectively. Regarding the ODSIS, our results are similar to those obtained by Bentley and colleagues (2014) and Ito et al. (2015), who recommended a cut-off point of 8 (82% correct classifications) and 11 (82.3% of correct classifications), respectively. What these different cut-off scores indicate is that validation and replication studies, such as the present, are important to increase the reliability of findings across settings and countries.

The Cronbach alphas of the two instruments (.87 for OASIS and .94 for the ODSIS) demonstrate excellent internal consistency. These scores fall within the range (.80 $\leq \alpha \leq$.96) observed in the literature (Bentley et al., 2014; Bragdon et al., 2016; Campbell-Sills et al., 2008; Hermans et al., 2014; Ito et al., 2014; Moore et al., 2015; Norman et al., 2013; Norman et al., 2011; Norman et al., 2006), suggesting that the items were well validated and adequately understood by this study participants.

Results on the convergent and discriminant validity of the OASIS and the ODSIS were also satisfactory. The correlations between both measures and the remaining instruments were significant and in the expected directions. At this stage, it is important to note that, contrary to our expectations, the OASIS did not correlate stronger with a measure of anxiety (BAI) than with the measure of depression (BDI-II). This result may be due to high rates of comorbidity between these two constructs (Brown & Barlow, 2009). However, the correlations between the ODSIS and the BDI-II and the BAI were consistent with our expectations (i.e., stronger for the BDI-II), so another possibility is that people with a primary diagnosis of an anxiety disorder perhaps did not meet criteria for a diagnosis of depression, but presented high scores on depression symptoms. In fact, comorbitidy of symptoms in our sample was evident. For instance, note that the 209 participants (61.7% of the sample) who presented moderate or severe scores on the BAI, also presented moderate or severe scores on the BDI-II.

In the evaluation of sensitivity to therapeutic change, both changes in the OASIS and the ODSIS moderately correlated with changes with the BAI and the BDI-II, respectively. These findings suggest that both instruments can capture changes in anxious and depressive symptoms after cognitive-behavioral treatment. This analysis was also performed in the study of Norman and colleagues (2013) for the OASIS, with a sample of women who had been victims of partner violence. Their results also showed positive correlations with the reference measures, namely the Anxiety subscale of the Brief Symptom Inventory (BSI-18; Derogatis & Melisaratos, 1983) and the Clinician Administered Post-Traumatic Stress Disorder Scale for DSM-IV (CAPS; Blake et al., 1995). They found that the change scores in the OASIS correlated highly with change scores in both the CAPS (r=.64, p<.01) and the BSI-Anxiety (r=.92, p<.001). The authors noted that future studies should examine sensitivity to change of the measures in a sample that also included men and higher diagnostic variability, both of which we have addressed in the present study.

The present study is not without limitations. First, only a subsample of participants provided data for all assessment points, so the sensitivity to the apeutic change could only be calculated in a subset of participants. It would be interesting to replicate these analyses using a larger sample. Another limitation is that the distribution of the sex was not homogeneous, as women were more frequently represented in the sample. While this might have influenced the present investigation results, it should be noted that the prevalence of women with ED is significantly higher than that of men (WHO, 2017), so the study findings are actually representative of the distribution of EDs in the general population and therefore might be useful for general practice. Another shortcoming was that inter-rater reliability of diagnoses could not be investigated. This had implications, as the sensitivity to change could not be examined against diagnoses based on structured clinical interviews due to the inability to test their reliability. In addition, we were not able to analyze the test-retest reliability because all participants were enrolled in a research clinical trial about the efficacy and feasibility of the UP for transdiagnostic treatment of ED, which would affect the stability of scores (Blinded). It is also important to note that the cut-off points for the OASIS and the ODSIS were calculated according to the established cut-offs for the BAI and BDI-II, as opposed to using a sample of healthy controls and instead of using a clinical-based criterion. While the BAI and the BDI-II are well-established screening tools, the possibility that these measures introduce some bias in diagnosis should not be ignored and the combination of self-reports and clinical-based criteria should be addressed in future research. Finally, the cut-offs proposed in this study are based on a sample of people seeking treatment for ED in a specialized mental care in Spain and included patients with a primary ED diagnosis, so the findings may not be generalizable to other settings.

Conclusions

This study adds to the growing literature supporting the validity of the OASIS and ODSIS as screening tools for anxiety and mood disorders. It also provides data about their sensitivity to change after a psychological treatment, which makes them useful tools to be included in routine care to rapidly assess the effectiveness of interventions. Additionally, this investigation provides further evidence for the utility of both scales in a new country, namely, Spain.

The validation of the OASIS and the ODSIS in Spanish offers large potential benefits to public health services. First, the OASIS and ODSIS are easily administered screening tools with adequate psychometric properties, allowing practitioners to detect, in just three minutes, individuals who present anxious or depressive symptomatology. In addition, because these scales do not require specialized training for their application, the OASIS and ODSIS may help family physicians to make important decisions, such as referring a person who needs specialized care. In to the case of mental health professionals working in specialized care, these scales can be useful to monitor the evolution of anxiety and depression symptoms over the course treatment in a quickly and easy way.

In conclusion, the OASIS and ODSIS have the potential to help improve patients' mental health on a scalable level, facilitating specialized, efficient, and prompt care for those who need it.

References

American Psychiatric Association., 1994. Diagnostic and statistical manual of mental disorders, fourth ed. Author, Washington, DC.

American Psychiatric Association., 2013. Manual diagnóstico y estadístico de los trastornos mentales, quinta ed. Madrid.[Diagnostic and statistical manual of mental disorders-text revision, fifth ed. Author, Washington, DC.

Barlow, D. H., Farchione, T. J., Bullis, J. R., Gallagher, M. W., Murray-Latin, H., Sauer Zavala, S., ...Ametaj, A., 2017. The unified protocol for transdiagnostic treatment of emotional disorders compared with diagnosis-specific protocols for anxiety disorders: A randomized clinical trial. JAMA.74, 875-884.

Beck, A. T., Steer, R., 1993. Beck Anxiety Inventory manual. Psychological Corporation, San Antonio, TX.

Beck, A. T., Steer, R. A., Brown, G.K., 1996. Manual for the Beck Depression Inventory-II.

Psychological Corporation, San Antonio, TX.

Bentley, K. H., Gallagher, M. W., Carl, J. R., Barlow, D. H., 2014.Development and validation of the Overall Depression Severity and Impairment Scale. Psychol Assess. 26, 815–830. https://doi.org/10.1037/a0036216

Blake D.D., Weathers F.W., Nagy L.M., Kaloupek D.G., Gusman F.D., Charney D.S., Keane, T.M., 1995. The development of a clinician-administered PTSD scale. J. Trauma Stress.

8, 75-90.

Botella, C., Ballester, R., 1997. Trastorno de pánico, evaluación y tratamiento [Panic disorder, evaluation and treatment]. Martínez Roca, Barcelona.

Bragdon, L. B., Diefenbach, G. J., Hannan, S., Tolin, D. F., 2016. Psychometric properties of the Overall Anxiety Severity and Impairment Scale (OASIS) among psychiatric outpatients. J.

Affect. Disord. 201, 112–115. https://doi.org/10.1016/j.jad.2016.05.005

Brown, T. A., Barlow, D. H., 2009. A proposal for a dimensional classification system based on the shared features of the DSM-IV Anxiety and mood disorders: Implications for assessment and treatment. Psychol Assess. 21, 256–271. https://doi.org/10.1037/a0016608

Brown, T. A., Di Nardo, P. A., Barlow, D. H., 1994. Anxiety Disorder Interview Schedule for DSM-IV (ADIS-IV). Adult and Lifetime version. Clinical Manual. Psychological Corporation, San Antonio, TX.

Brown, T.A., Di Nardo, P.A., Lehman, C.L., Campbell, L.A., 2001. Reliability of DSM-IV anxiety and mood disorders: Implications for the classification of emotional disorders. J.

Abnorm. Psychol. 110, 49-58. https://doi.org/10.1037//0021-843X.110.1.49

Campbell-Sills, L., Norman, S. B., Craske, M. G., Sulivan, G., Lang, A. J., Chavira, D. A., Stein, M.B., 2009. Validation of a brief measure of anxiety-related severity and impairment: the Overall Anxiety Sensitivity and Impairment Scale (OASIS). J. Affect. Disord. 112, 92–101. https://doi.org/10.1016/j.jad.2008.03.014

Cano-Vindel, A., Salguero, J. M., Mae Wood, C., Dongil, E., Latorre Postigo, J. M., 2012.La depresión en atención primaria: prevalencia, diagnóstico y tratamiento [Depression in primary care: Prevalence, diagnosis, and treatment]. Papeles del Psicólogo. 33, 2–11.

Castro-Rodríguez, J. I., Olariu, E., Garnier-Lacueva, C., Martín-López, L. M., Pérez-Solà, V., ... Insa, I., 2015. Diagnostic accuracy and adequacy of treatment of depressive and anxiety disorders: A comparison of primary care and specialized care patients. J. Affect. Disord.

172, 462–471. https://doi.org/10.1016/j.jad.2014.10.020

Chueca, I. M., María, M. P., Bayo, I. G., Durán, E. C., Vila, M. S., Blanco, A. S., 2003. ¿Cómo derivamos a salud mental desde atención primaria? Atención Primaria. 32(9), 524-530. https://doi.org/10.1016/S0212-6567(03)70782-3

Costa, P. T., McCrae, R. R., 1999. Revised NEO Personality Inventory (NEO-PI-R) and NEO Five-Factor Inventory (NEO-FFI). TEA Ediciones, Madrid.

Derogatis, L.R., Melisaratos, N., 1983. The brief symptom inventory: an introductory report.

Psychol Med. 13, 595-605.

Deveugele, M., Derese, A., Brink-Muinen, A., Bensing, J., De Maeseneer, J., 2002.

Consultation length in general practice: Cross-sectional study in six European countries.

Br. Med. J. 325, 472–477. https://doi.org/10.1136/bmj.325.7362.472

Diedenhofen, B., Musch, J. 2015. cocor: A Comprehensive Solution for the Statistical Comparison of Correlations. PLoS ONE. 10(4): e0121945. https://doi.org/10.1371/journal.pone.0121945

Echeburúa, E., Corral, P., Fernández-Montalvo, J., 2000. Maladjustment Inventory (MI)

Psychometric properties in clinical contexts. Análisis y Modif. Conduct. 26, 325–340.

Ferrans, C., Powers, M., 1985. Quality of Life Index: Development and psychometric properties. Adv. Nurs. Sci. 8, 15-24.

First, M. B., Spitzer, R. L., Gibbon, M., Williams, J. B., 1997. User's guide for the Structured Clinical Interview for DSM-IV, Axis I Disorders (SCID-I). American Psychiatric

Association, Washington, DC.

First, M. B., Spitzer, R. L., Williams, J. B., Gibbon, M., 1999. Entrevista Clínica Estructurada para los Trastornos del Eje I del DSM-IV, Versión Clínica (SCID-I-VC) [Structured clinical Interview for Axis I disorders of the DSM-IV, clinical version]. Masson, Barcelona.

González, N. L., Sarriés, A. G., de Jalón Aramayo, E. G., Urra, E. L., 2008. Concordancia en el diagnóstico entre atención primaria y salud mental [Diagnosis agreement between primary care and mental health]. Atención primaria. 40(6), 285-289. https://doi.org/10.1157/13123680

Hermans, M., Korrelboom, K., Visser, S., 2014. A Dutch version of the Overall Anxiety

Severity and Impairment Scale (OASIS): Psychometric properties and validation. J.

Affect. Disord. 172, 127–132. https://doi.org/10.1016/j.jad.2014.09.033

Hu, L., Bentler, P. M., 1999. Cutoff criteria for fit indexes in covariance structure analysis:

Conventional criteria versus new alternatives. Struc. Equ. Model. A Multidiscip. J. 6(1),

1-55. http://doi.org/10.1080/10705519909540118

IBM Corp, n.d. IBM SPSS Statistics for Windows, Version 22.0.

Ito, M., Bentley, K. H., Oe, Y., Nakajima, S., Fujisato, H., Kato, N., ...Barlow, D. H., 2015. Assessing depression-related severity and functional impairment (warning): The Overall Depression Severity and Impairment Scale (ODSIS).PLoS ONE. 10, 1–14.

https://doi.org/10.1371/journal.pone.0122969

Ito, M., Oe, Y., Kato, N., Nakajima, S., Fujisato, H., Miyamae, M., ... Norman, S. B., 2014.

Validity and clinical interpretability of the Overall Anxiety Severity and Impairment Scale

(OASIS). J. Affect. Disord. 170, 217–224. https://doi.org/10.1016/j.jad.2014.08.045

Kroenke, K., Spitzer, R. L., Williams, J. B., 2001. The PHQ- 9: validity of a brief depression severity measure. J. Gen. Intern. Med. 16, 606-613. https://doi.org/10.1046/j.1525-

1497.2001.016009606.x

Latorre, J. M., López-Torres, J., Montañés, J., Parra, M., 2005. Percepción de la demanda y necesidad de formación de los médicos de atención primaria en salud mental [Perception of the demand and need for training of primary care physicians in mental health]. Atención

Primaria. 36, 85-92. https://doi.org/10.1157/13076608

Latorre Postigo, J. M., Navarro Bravo, B., Parra Delgado, M., Salguero, J. M., Mae Wood, C., Cano Vindel, A., 2012. Evaluación e intervención de los problemas de ansiedad y depresión en atención primaria: Un problema sin resolver [Assessment and intervention in anxiety and depression problems in primary care: An unsolved problem]. Rev. Clínica

Med. Fam. 5, 37-45. https://doi.org/10.4321/S1699-695X2012000100007

Mezzich, J. E., Ruipérez, M. A., Pérez, C., Yoon, G., Liu, J., Mahmud, S., 2000. The Spanish version of the Quality Of Life Index: Presentation and validation. J. Nerv. Ment. Dis. 188, 301–305.

Moore, S. A., Welch, S. S., Michonski, J., Poquiz, J., Osborne, T. L., Sayrs, J., Spanos, A.,

2015.Psychometric evaluation of the Overall Anxiety Severity and Impairment Scale (OASIS) in individuals seeking outpatient specialty treatment for anxiety-related disorders. J. Affect. Disord.175, 463–470. https://doi.org/10.1016/j.jad.2015.01.041 Muñoz-Navarro, R., Cano-Vindel, A., Ruiz-Rodríguez, P., Medrano, L. A., González-Blanch, C.,... Dongil-Collado, E., 2017. Modelo jerárquico de diagnóstico y derivación de los trastornos mentales comunes en centros de atención primaria. Una propuesta a partir del ensayo clínico PsicAP [Hierarchical model of diagnosis and referral of common mental disorders in primary care. A proposal from the Psic AP clinical trial]. Ansiedad y Estrés. 23,

124–129. https://doi.org/10.1016/j.anyes.2017.10.002

Muthen, L. K., Muthen, B. O., 1998. Mplus User's Guide, sixth ed. Muthén&Muthén, Los Angeles, CA.

Norman, S. B., Allard, C. B., Trim, R. S., Thorp, S. R., Behrooznia, M., Masino, T. T., Stein,

M. B., 2013. Psychometrics of the Overall Anxiety Severity and Impairment Scale

(OASIS) in a sample of women with and without trauma histories. Arch. Women's Ment.

Heal. 16, 123–129. https://doi.org/10.1007/s00737-012-0325-8

Norman, S. B., Campbell-Sills, L., Hitchcock, C. A., Sullivan, S., Rochlin, A., Wilkins, K. C., Stein, M. B., 2011. Psychometrics of a brief measure of anxiety to detect severity and impairment: The Overall Anxiety Severity and Impairment Scale (OASIS). J. Psychiatr. Res. 45, 262–268. https://doi.org/10.1016/j.jpsychires.2010.06.011

Norman, S. B., Cissell, S., Means- Christensen, A. J., Stein, M. B., 2006. Development and

validation of an Overall Anxiety Severity and Impairment Scale (OASIS). Depress

Anxiety. 23, 245–249. https://doi.org/10.1002/da20182

Osma, J., Castellano, C., Crespo, E., García-Palacios, A., 2015. The unified protocol for transdiagnostic treatment of emotional disorders in group format in a Spanish public mental health setting. Behav. Psychol. 23, 447-466.

Rogers, R., 2001. Handbook of diagnostic and structured interviewing. Guilford Press, New York, NY.

Ruiz-Rodríguez, P., Cano-Vindel, A., Navarro, R. M., Medrano, L., Moriana, J. A., Aguado, C. B.,...Grupo de Investigación PsicAP., 2017. Impacto económico y carga de los trastornos mentales comunes en España: Una revisión sistemática y crítica [Economic impact and burden of common mental disorders in Spain: A systematic and critical review]. Ansiedad y Estrés. 23, 118–123. https://doi.org/10.1016/j.anyes.2017.10.003

Sandín, B., Chorot, P., Lostao, L., Joiner, T. E., Santed, M. A., Valiente, R. M., 1999. PANAS Positive and Negative Affect Schedule: Factorial validation and transcultural convergence. Psicothema. 11, 37–51.

Sanz, J., 2014. Recomendaciones para la utilización de la adaptación española del Inventario de Ansiedad de Beck (BAI) en la práctica clínica [Recommendations for the use of the Spanish adaptation of the Beck Anxiety Inventory (BAI) in clinical practice]. Clínica y Salud. 25(1), 39-48. http://dx.doi.org/10.5093/cl2014a3

Sanz, J., García-Vera, M. P., Espinosa, R., Fortún, M., Vázquez, C., 2005. Adaptación española del Inventario para la Depresión de Beck-II (BDI-II): 3. Propiedades psicométricas en pacientes con trastornos psicológicos [Spanish adaptation of the Beck-II Depression Inventory (BDI-II): 3. Psychometric properties in patients with psychological disorders].

Clínica y Salud. 16, 121–142.

Sanz, J., García-Vera, M. P., Fortún, M., 2012. El Inventario de Ansiedad de Beck (BAI): propiedades psicométricas de la versión española en pacientes con trastornos psicológicos [The Beck Anxiety Inventory (BAI): Psychometric properties of the Spanish version in patients with psychological disorders]. Psicol. Conductual. 20, 563.

Sanz, J. Navarro, M. E., 2003. Propiedades psicométricas de una versión española del Inventario de Ansiedad de Beck (BAI) en estudiantes universitarios [Psychometric properties of a Spanish version of the Beck Anxiety Inventory (BAI) in college students].

Ansiedad y Estrés. 9, 59–84.

Sanz, J., Perdigón, A. L., Vázquez, C., 2003. The Spanish adaptation of the Beck Depression Inventory-II (BDI-II): 2. Psychometric properties in the general population. Clínica y Salud. 14, 249–280.

Shear, M. K., Rucci, P., Williams, J., Frank, E., Grochocinski, V., Bander Bilt, J., ... Wang, T., 2001. Reliability and validity of the Panic Disorder Severity Scale: Replication and extension. J. Psychiatry Res. 35, 293-296.

Spitzer, R. L., Kroenke, K., Williams, J. B., Löwe, B., 2006. A brief measure for assessing generalized anxiety disorder: the GAD-7. Arch. Intern. Med. 166, 1092-1097.

https://doi.org/doi:10.1001/archinte.166.10.1092

Watson, D., Clark, L. A., Tellegen, A., 1988. Development and validation of brief measures of positive and negative affect: The PANAS Scales. J. Pers. Soc. Psychol. 54, 1063–1070.

World Health Organization., 2017. Depression and other common mental disorders: Global health estimates. Author, Geneva, Switzerland.

Zweig, M. H., Campbell, G., 1993. Receiver-operating characteristic (ROC) plots: A fundamental evaluation tool in clinical medicine. Clin. Chem. 39, 561–577.

Conflict of Interest

The authors declare that they have no conflict of interest.

Ethical approval

This research was conducted with the approval of the ethics committee of all collaborating centers and all participants signed an informed consent prior to participation. All procedures performed in studies involving human participants were in accordance with the ethical standards of the institutional and/or national research committee and with the 1964 Helsinki declaration and its later amendments or comparable ethical standards. This article does not contain any studies with animals performed by any of the authors.

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Contributors

All persons who met authorship criteria are listed as authors. All authors certify that they have participated sufficiently in the work to take public responsibility for the content, including participation in the concept, design, analysis, writing, or revision of the manuscript. All authors contributed to and approved the final manuscript.

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Appendix 1

Both scales have been reprinted from Barlow, D. H., Sauer-Zavala, S., Farchione, T. J., Murray, H., Ellard, K., Bullis, J.,...Cassiello-Robbins, C. (2018). Unified Protocol for Transdiagnostic Treatment of Emotional Disorders (2nded.) [Protocolo Unificado para el Tratamiento Transdiagnóstico de los Trastornos Emocionales (2.ª ed.), traducida por Jorge Osma y Elena Crespo)]. Madrid: Alianza Editorial, 2019.

Escala General de Gravedad e Interferencia de Ansiedad

Los siguientes ítems preguntan sobre ansiedad. Para cada ítem, rodee el número que mejor describa su experiencia durante la última semana.

1. Durante la última semana, ¿con qué frecuencia te has sentido ansioso?

- = No me sentí ansioso durante la última semana.
- = Ansiedad infrecuente. Me sentí ansioso en algunos momentos.
- Ansiedad ocasional. La mitad del tiempo me sentí ansioso y la otra mitad no. Me costó relajarme.
- Ansiedad frecuente. Me sentí ansioso la mayor parte del tiempo. Me resultó muy difícil relajarme.
 - Ansiedad constante. Me sentí ansioso todo el tiempo y nunca llegué a relajarme.

2. Durante la última semana, cuando te sentiste ansioso, ¿cómo de intensa o grave fue tu ansiedad?

- Poco o nada. La ansiedad estuvo ausente o casi no la noté.
- = Leve. La ansiedad fue de baja intensidad. Pude relajarme cuando lo intenté. Los síntomas físicos fueron solo un poco molestos.
- = Moderada. La ansiedad me generó malestar en algunos momentos. Me resultó dificil relajarme o concentrarme, pero pude hacerlo cuando lo intenté. Los síntomas físicos fueron molestos.
- = Severa. La ansiedad fue intensa la mayor parte del tiempo. Me resultó muy difícil relajarme o concentrarme en cualquier otra cosa. Los síntomas físicos fueron enormemente molestos.
- = Extrema. La ansiedad me sobrepasó. Me fue totalmente imposible relajarme. Los síntomas físicos fueron insoportables.

3. Durante la última semana, ¿con qué frecuencia evitaste situaciones, lugares, objetos o actividades debido a tu ansiedad o miedo?

- Ninguna. No evité lugares, situaciones, actividades o cosas por miedo.
- | = Infrecuente. Evité algunas cosas de vez en cuando, pero por lo general me enfrenté a las situaciones u objetos. Mi estilo de vida no se vio afectado.
- — Ocasional. Tuve algo de miedo a ciertas situaciones, lugares u objetos, pero todavía pudo manejarlos. Mi estilo de vida sufrió pocos cambios. Siempre o casi siempre evité las cosas que me dan miedo si estaba solo, pero las pude manejar si alguien venía conmigo.
- = Frecuente. Tuve bastante miedo y realmente intenté evitar las cosas que me asustan. He hecho cambios significativos en mi estilo de vida para evitar objetos, situaciones, actividades o lugares.
- Todo el tiempo. Evitar objetos, situaciones, actividades o lugares ha ocupado gran parte de mi vida. Mi estilo de vida se ha visto enormemente afectado y ya no hago cosas con las que solía disfrutar.

- 4. Durante la última semana, ¿en qué medida ha interferido la ansiedad en tu capacidad para hacer las cosas que necesitabas hacer en el trabajo, el colegio o en tu hogar?
 - Nada. La ansiedad no interfirió en mi trabajo/hogar/colegio.
 - = Leve. La ansiedad me causó algo de interferencia en mi trabajo/hogar/colegio. Las cosas eran más difíciles, pero pude realizar todo lo que necesitaba hacer.
 - = Moderada. La ansiedad definitivamente interfirió en mis tareas. He podido realizar la mayoría de las cosas, pero solo algunas las he hecho tan bien como en el pasado.
 - = Severa. La ansiedad verdaderamente ha cambiado mi capacidad para hacer las cosas. Algunas cosas las he podido realizar, pero otras no. Mi rendimiento se ha visto definitivamente afectado.
 - Extrema. La ansiedad ha llegado a ser incapacitante. He sido incapaz de completar mis tareas y he tenido que irme del colegio, he dejado o me han despedido de mi trabajo o he sido incapaz de completar las tareas del hogar y he sufrido consecuencias como desalojos, cobradores de facturas, etc.
- 5. Durante la última semana, ¿en qué medida ha interferido la ansiedad en tu vida social y en tus relaciones?
 - = Nada. La ansiedad no interfirió en mis relaciones.
 - Leve. La ansiedad apenas interfirió en mis relaciones. Algunas de mis amistades y otras relaciones se han visto afectadas, pero en conjunto mi vida social sigue siendo satisfactoria.
 - = Moderada. La ansiedad interfirió algo en mi vida social, pero sigo teniendo algunas relaciones cercanas. No paso tanto tiempo con otros como en el pasado, pero sigo teniendo relaciones sociales algunas veces.
 - = Severa. Mis amistades y otras relaciones se han visto muy afectadas a causa de mi ansiedad. No disfruto de las actividades sociales. Tengo muy pocas relaciones sociales.
 - Extrema. La ansiedad ha alterado completamente mis actividades sociales. Todas mis relaciones se han visto afectadas o han finalizado. Mi vida familiar es extremadamente tensa.

P	untuación	total:	

Escala General de Gravedad e Interferencia de Depresión

Los siguientes ítems preguntan sobre depresión. Para cada ítem, rodee el número que mejor describa su experiencia durante la última semana.

- 1. Durante la última semana, ¿con qué frecuencia te has sentido deprimido?
 - No me sentí deprimido durante la última semana.
 - Depresión infrecuente. Me sentí deprimido en algunos momentos.
 - Depresión ocasional. La mitad del tiempo me sentí deprimido y la otra mitad no.
 - Depresión frecuente. Me sentí deprimido la mayor parte del tiempo.
 - = Depresión constante. Me sentí deprimido todo el tiempo.
- 2. Durante la última semana, cuando te sentiste deprimido, ¿cómo de intensa o grave fue tu depresión?
 - = Poco o nada. La depresión estuvo ausente o casi no la noté.

- = Leve. La depresión fue de baja intensidad.
- = Moderada. La depresión me generó malestar en algunos momentos.
- = Severa. La depresión fue intensa la mayor parte del tiempo.
- Extrema. La depresión me sobrepasó.

3. Durante la última semana, ¿con qué frecuencia tuviste dificultad para realizar o interesarte en actividades que normalmente disfrutas debido a tu depresión?

- = Ninguna. No tuve dificultades para realizar o interesarme en actividades que normalmente disfruto debido a la depresión.
- = Infrecuente. Algunas veces tuve dificultades para realizar actividades o interesarme en actividades que normalmente disfruto, debido a la depresión. Mi estilo de vida no se vio afectado.
- — Ocasional. Tuve algunas dificultades para realizar actividades o interesarme en actividades que normalmente disfruto, debido a la depresión. Mi estilo de vida sufrió pocos cambios.
- Frecuente. Tuve bastantes dificultades para realizar actividades o interesarme en actividades que normalmente disfruto, debido a la depresión. He realizado cambios significativos en mi estilo de vida por no poder interesarme en actividades que solía disfrutar.
- Todo el tiempo. No he podido participar o interesarme en actividades que normalmente disfruto, debido a la depresión. Mi estilo de vida se ha visto enormemente afectado y ya no hago cosas que solía disfrutar.

4. Durante la última semana, ¿en qué medida ha interferido la depresión en tu capacidad para hacer las cosas que necesitabas hacer en el trabajo, el colegio o en tu hogar?

- = Nada. La depresión no interfirió en mi trabajo/hogar/colegio.
- = Leve. La depresión me causó algo de interferencia en mi trabajo/hogar/colegio. Las cosas fueron más difíciles, pero pude realizar todo lo que necesitaba hacer.
- = Moderada. La depresión definitivamente interfirió en mis tareas. He podido realizar la mayoría de las cosas, pero solo algunas las he hecho tan bien como en el pasado.
- = Severa. La depresión verdaderamente ha interferido en mis tareas. Algunas tareas las he podido realizar, pero muchas otras no. Mi rendimiento se ha visto definitivamente afectado.
- Extrema. La depresión ha llegado a ser incapacitante. He sido incapaz de completar mis tareas y he tenido que irme del colegio, he dejado o me han despedido de mi trabajo o he sido incapaz de completar las tareas del hogar y he sufrido consecuencias como desalojos, cobradores de facturas, etc.

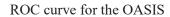
5. Durante la última semana, ¿en qué medida ha interferido la depresión en tu vida social y en tus relaciones?

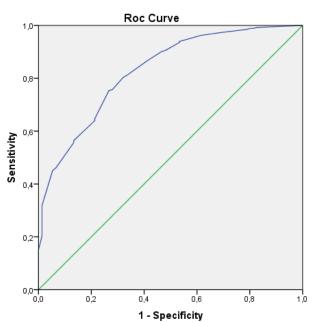
- Nada. La depresión no interfirió en mis relaciones.
- = Leve. La depresión apenas interfirió en mis relaciones. Algunas de mis amistades y otras relaciones se han visto afectadas, pero en conjunto mi vida social sigue siendo satisfactoria.
- = Moderada. La depresión ha interferido algo en mi vida social, pero sigo teniendo algunas relaciones cercanas. No paso tanto tiempo con otros como en el pasado, pero sigo manteniendo relaciones sociales algunas veces.
- = Severa. Mis amistades y otras relaciones se han visto muy afectadas a causa de mi depresión. No disfruto de las actividades sociales. Tengo muy pocas relaciones sociales.

• Extrema. La depresión ha alterado completamente mis actividades sociales. Todas mis relaciones se han visto afectadas o han finalizado. Mi vida familiar es extremadamente tensa.

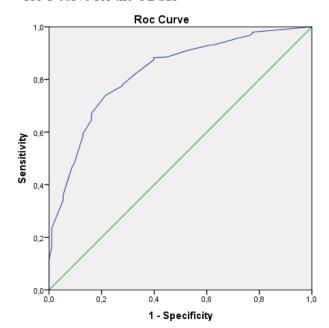
Puntuación total:







ROC curve for the ODSIS



Socio-demographic Characteristics of Participants (N = 339)

Table 1

Socio-demographic Characteristics of Part	Participants with one	Participants with	χ^2	p
	evaluation	two evaluations		
	n (%)	n (%)		
Educational level	120 (35.4)	219 (64.6)	0.991	.319
Less than 12 years of education received	51 (42.5)	81 (37.0)		
No studies	2 (1.7)	5 (2.3)		
Primary studies	24 (20.0)	38 (17.4)		
Secondary studies	25 (20.8)	38 (17.4)		
More than 12 years of education received	69 (57.5)	138 (63.0)		y
High school	9 (7.5)	28 (12.8)		
Vocational training	23 (19,2)	62 (28,3)	, 7	
University studies	37 (30.8)	48 (21.9)		
Marital status	120 (35.4)	219 (64.6)	3.554	.059
Married/living w partner	48 (40.0)	111 (50.7)		
Not Married/not living w partner	72 (60.0)	108 (49.3)		
Single	51 (42.5)	67 (30.6)		
Separated/ Divorced	17 (14.2)	36 (16.4)		
Widowed	4 (3.3)	5 (2.3)		
Job status	120 (35.4)	219 (64.6)	1.968	.161
Working	56(46.7)	85 (38.8)		
Non-working	64 (53.3)	134 (61.2)		
Unemployed	26 (21.7)	52 (23.7)		
Sick leave	15 (12.5)	44 (20.1)		
Student	14 (11.7)	12 (5.5)		
Home-maker	3 (2.5)	20 (9.1)		
Retiree	6 (5.0)	6 (2.7)		
Primary diagnosis	120 (35.4)	219 (64.6)	3.984	.136
Anxiety disorders	65 (54.2)	102 (46.6)		
Generalized anxiety disorder	17 (14.2)	26 (11.9)		
Panic disorder with agoraphobia	22 (18.3)	13 (5.9)		
Non-specific anxiety disorder	4 (3.3)	15 (6.8)		
Panic disorder without agoraphobia	5 (4.2)	12 (5.5)		
Agoraphobia	7 (5.8)	8 (3.7)		
Obsessive-compulsive disorder	2 (1.7)	10 (4.6)		
Social phobia	5 (4.2)	5 (2.3)		
Posttraumatic stress disorder	1 (.8)	6 (2.7)		
Hypochondria	2 (1.7)	4 (1.8)		

Specific phobia food disorders Major depressive disorder	$\Omega \setminus \Omega $	0 (1 1)		
	0 (.0)	3 (1.4)		
Major depressive disorder	35 (29.2)	60 (27.4)		
	22 (18.3)	40 (18.3)		
Dysthymia	12 (10.0)	11 (5.9)		
Unspecified mood disorder	1 (.8)	9 (4.1)		
fixed disorders	20 (16.7)	57 (26.9)		
djustment disorder	20 (16.7)	57 (26.9)		
condary diagnosis	37 (10.9)	62 (18.3)	3.159	.206
nxiety disorders	25 (20.8)	40 (18.3)		
Non-specific anxiety disorder	3 (2.5)	11 (5.0)		
Obsessive-compulsive disorder	5 (4.2)	5 (2.3)		
Generalized anxiety disorder	2 (1.7)	6 (2.7)) 7
Social phobia	4 (3.3)	3 (1.4)		
Panic disorder without agoraphobia	2 (1.7)	4 (1.8)		
Agoraphobia	4 (3.3)	2 (.9)		
Specific phobia	2 (1.7)	3 (1.4))	
Hypochondria	0 (0.0)	4 (1.8)		
Panic disorder with agoraphobia	1 (.8)	2 (.9)		
Posttraumatic stress disorder	2 (1.7)	0 (.0)		
lood disorders	8 (6.7)	20 (9.1)		
Major depressive disorder	7 (5.8)	12 (5.5)		
Dysthymia	1 (.8)	8 (3.7)		
lixed disorders	4 (.3)	2 (.9)		
Adjustment disorder	3 (2.5)	2 (.9)		
Somatoform disorder	1 (.8)	0 (.0)		

Table 2
Correlations between the OASIS and the ODSIS and Measures of Convergent and Discriminant Validity and internal consistency estimates

	M (SD)	Min - Max	OASIS	ODSIS	BAI	BDI-II	NEO-FFI N	NEO-FFI E	QLI	SM	PANAS POS	PANAS NEG
OASIS	10.45 (4.49)	0 - 20	(.87)	.69**	.57**	.60**	.43**	30**	57**	.57**	34**	.49**
ODSIS	9.88 (5.14)	0 - 20		(.94)	.42**	.68**	.43**	33**	64**	.58**	48**	.42**
BAI	26.81 (13.36)	0 - 63			(.93)	.53**	.42**	20**	43**	.41**	20**	.50**
BDI-II	27.60 (12.12)	0 - 63				(.92)	.50**	40**	69**	.63**	49**	.53**
NEO-FFI N	31.64 (7.62)	0 - 48					(.76)	34**	45**	.43**	33**	.56**
NEO-FFI E	22.70 (8.01)	0 - 48						(.81)	.48**	35**	.46**	19**
QLI	46.05 (16.06)	0 - 100							(.87)	64**	.51**	40**
SM	18.09 (6.33)	0 - 30				•				(.83)	36**	.41**
PANAS-POS	20.72 (6.90)	10 - 50				•		1			(.89)	16**
PANAS-NEG	28.49 (8.72)	10 - 50										(.89)

Note: OASIS: Overall Anxiety Severity and Impairment Scale; ODSIS: Overall Depression Severity and Impairment Scale; BAI: Beck Anxiety Inventory; BDI-II: Beck Depression Inventory; NEO-FFI N: NEO-Five-Factor Personality Inventory Neuroticism; NEO-FFI E: NEO-Five-Factor Personality Inventory Extraversion; QLI: Quality of Life Index; SM: Scale of Maladjustment; PANAS POS: Positive and Negative Affect Scale- Positive; PANAS NEG: Positive and Negative Affect Scale- Negative. Minimum and maximum scores correspond to direct scale scores. Scale reliability corresponds to the Cronbach's alpha coefficient and is presented in parentheses. *p < .01 **p < .001

Table 3

Factor Loadings with the 5-Factor Model without Correlating the Error Variance for the OASIS

Item	Description Factor loading			
		OASIS	ODSIS	
1	Frequency of symptoms	.62	.82	
2	Intensity of symptoms	.68	.81	
3	Degree of avoidance/ loss of interest	.68	.90	
4	Interference in work, school life	.89	.88	
5	Interference in social life	.86	.86	

In all cases, p < .001.